

## ASYMPTOTIC BEHAVIOUR OF THE INTEGRAL OF A FUNCTION ON THE LEVEL SET OF A MIXING RANDOM FIELD\*

BY

ILEANA IRIBARREN (CARACAS)

*Abstract.* Let  $X = \{X(t): t \in \mathbb{R}^2\}$  be a centered stationary real random field with a.s. differentiable paths. Let  $T$  be a rectangle of  $\mathbb{R}^2$  and let  $F(f, T)$  denote the integral of the continuous function  $f$  over a level curve  $\mathcal{C}_x$  of  $X$  for a fixed level  $x$ , observed in  $T$ . We show that if a field  $X$  satisfies some mixing condition, then  $F(f, T)$ , adequately normalized, converges weakly to the Wiener process indexed in  $T$ . The limit variance has a precise expression in the Gaussian case and \*-mixing case. A geometrical lemma shows cases where the higher order moments of  $F(f, T)$  are finite.

### 1. INTRODUCTION.

In what follows,  $X = \{X(t): t \in \mathbb{R}^2\}$  denotes an a.s. differentiable centered stationary random field.  $X$  is said to be *isotropic* if, given any  $k \in \mathbb{N}$  and  $t_1, t_2, \dots, t_k \in \mathbb{R}^2$ , the joint laws of  $(X(t_1), X(t_2), \dots, X(t_k))$  and  $(X(Qt_1), X(Qt_2), \dots, X(Qt_k))$  are the same when  $Q$  is any isometry in  $\mathbb{R}^2$ .  $X$  is said to be *affine*, when it is equal in law to  $\{Y(At): t \in \mathbb{R}^2\}$ , where  $Y$  is isotropic and  $A: \mathbb{R}^2 \rightarrow \mathbb{R}^2$  is a linear self adjoint transformation. The angle  $\theta_0$  defining the eigenvectors directions  $(\cos \theta_0, \sin \theta_0)$  and  $(-\sin \theta_0, \cos \theta_0)$  and the respective eigenvalues  $\lambda_1, \lambda_2$  specify the affinity  $A$ . There is no loss of generality in assuming  $\lambda_1 \geq \lambda_2 > 0$ ,  $\lambda_1 \lambda_2 = 1$ . Cabaña [2] has proposed estimators of the *affinity parameters*  $k = (1 - \lambda_1^2 / \lambda_2^2)^{1/2}$  and  $\theta_0$  based on the shape of the level curve of  $X$ , corresponding to a given level  $x$ .

The *gradient process*  $\dot{X}(t) = \|\dot{X}(t)\| (\cos \Theta(t), \sin \Theta(t))$  is determined by the two real stationary processes  $\|\dot{X}\|$  and  $\Theta$ . If  $X$  has a.s. differentiable Jacobian and  $T$  is an open rectangle of  $\mathbb{R}^2$ , the set  $\mathcal{C}_x = \{t \in T: X(t) = x, \dot{X}(t) \neq 0\}$ , for a fixed  $x$ , is a.s. a  $C^1$  one-dimensional manifold, as it results from applying the implicit Function Theorem. For  $f: (-\pi, \pi] \rightarrow \mathbb{R}$

---

\* Some of the results presented here are part of the author's Doctoral dissertation at Universidad Central de Venezuela.

continuous and bounded we write

$$F(f, T) = |T|^{-1} \int_{\mathcal{C}_x} f(\Theta(t)) dS(t),$$

where  $|T|$  is the area of  $T$ ,  $dS$  is the surface measure of  $\mathcal{C}_x$ , and the integral is defined for every  $\omega \in \Omega$  as the usual integral over a differentiable manifold [3].  $F(f, T)$  is a random variable with respect to the  $\sigma$ -algebra generated by  $\{X(t): t \in T\}$ . We write

$$\mathcal{L}(T) = F(1, T), \quad \mathcal{C}(T) = F(\cos 2\theta, T), \quad \mathcal{S}(T) = F(\sin 2\theta, T).$$

Cabaña [3] and Wschebor [10] have proved by different methods the formulae, known as *Rice's Formulae*, for the moments of the variable  $F(f, T)$  (Theorem 3). Making use of these formulae we obtain

$$\tan 2\theta_0 = \mathcal{S}/\mathcal{C}, \quad g(k) = \frac{\sqrt{(\mathcal{E}\mathcal{C})^2 + (\mathcal{E}\mathcal{S})^2}}{\mathcal{E}\mathcal{L}},$$

where  $g$  is a certain function of the parameter  $k$ . These equations led Cabaña to propose the following estimators for the affinity parameters  $\theta_0$  and  $k$ :

$$\theta_0 = \frac{1}{2} \arg(\mathcal{C}(\lambda T), \mathcal{S}(\lambda T)), \quad k = g^{-1} \left( \frac{\sqrt{\mathcal{C}^2(\lambda T) + \mathcal{S}^2(\lambda T)}}{\mathcal{L}(\lambda T)} \right),$$

where  $\lambda T$  is a dilation of  $T$ ,  $\theta_0$  and  $k$  are consistent and asymptotically Gaussian under adequated mixing conditions. Cabaña [2] has proved this assertion when  $X$  is  $\delta$ -dependent, and he has computed the limit variance. The aim of the present paper is to prove the asymptotical normality of  $F(f, T)$  under less strict conditions of dependence. We prove, making use of a Functional Central Limit Theorem for strong mixing fields, that adequately normalized  $F(f, T)$  converges weakly to a Wiener process indexed by the rectangles of  $\mathbb{R}^2$ . Generally, the Central Limit Theorem for mixing variables has the inconvenient that the variance is impossible to compute, but here we compute the limit variance in the Gaussian and \*-mixing cases and show that this variance has the same expression that in the  $\delta$ -dependent case (see [2]).

Since the existence of higher order moments plays a fundamental role, we improve Wschebor's results [11], giving conditions under which the  $p$ -th order moment of  $F(f, T)$  is finite.

## 2. FRAMEWORK

We begin with

**Definition 1.** Let  $\xi = \{\xi(t): t \in \mathbb{R}^d\}$  be a stationary random field. We say that  $\xi$  is *strongly mixing with coefficient  $\alpha$* , or  *$\alpha$ -mixing*, if for any Borelian

sets  $U$  and  $V$

$$\sup \{ |P(A \cap B) - P(A)P(B)| : A \in \sigma(U), B \in \sigma(V) \} \leq \alpha(p(U, V))$$

and  $\alpha(p) \rightarrow 0, p \rightarrow \infty$ . Here  $p(U, V)$  is the Euclidean distance and  $\sigma(U) = \sigma \{ \xi(t) : t \in U \}$ . In the same manner, we say  $\xi$  satisfies a  $\ast$ -mixing condition with coefficient  $\psi$  if

$$\sup \left\{ \left| \frac{P(A \cap B)}{P(A)P(B)} - 1 \right| : A \in \sigma(U), B \in \sigma(V) \right\} \leq \psi(p(U, V))$$

and  $\psi(p) \rightarrow 0, p \rightarrow \infty$ .

Let  $\xi = \{ \xi(t) : t \in \mathbb{Z}^d \}$  be a discrete strong mixing stationary centered real random field and

$$S_n = \sum_{1 \leq j \leq n} \xi_j,$$

where  $j \leq n$  means  $j_1 \leq n_1, j_2 \leq n_2, \dots, j_d \leq n_d$ . We prove

LEMMA 1. If  $E|\xi_0|^{2+\delta} < \infty$  and the mixing satisfies

$$\sum_{r=1}^{\infty} r^{d-1} \alpha^{\delta/(2+\delta)}(r) < \infty,$$

then

$$(i) \quad \sum_{j \in \mathbb{Z}^d} E(\xi_0 \xi_j) < \infty,$$

$$(ii) \quad |n|^{-1} ES_n^2 \rightarrow \sum_{j \in \mathbb{Z}^d} E(\xi_0 \xi_j) = \sigma^2 \quad \text{when } n \rightarrow \infty,$$

where  $n \rightarrow \infty$  means  $\min_{1 \leq j \leq d} n_j \rightarrow \infty$  and  $|n| = n_1 n_2 \dots n_d$ .

Proof. Lemma 3 of Billingsley [1], p. 172, and the inequality of moments for strong mixing fields [7] give the proof.

Let  $T^d$  be a Cartesian product of  $[0, 1] \subset \mathbb{R}$ . We will denote by  $C_d$  the set of continuous functions on  $T^d$  provided with the uniform metric, and  $D_d$  is the Skorohod function space on  $T^d$ . A subset  $B$  of  $T^d$ ,  $B = \{u = (u_1, \dots, u_d) : s_j \leq u_j \leq t_j\}$  is called a block, and increment  $\xi(B)$  of  $\xi$  around a block  $B$  is given by

$$\xi(B) = \sum_{\varepsilon \in \{0,1\}^d} (-1)^{d-|\varepsilon|} \xi(s + \varepsilon(t-s)),$$

where  $|\varepsilon| = \varepsilon_1 + \dots + \varepsilon_d$  if  $\varepsilon = (\varepsilon_1, \dots, \varepsilon_d)$ .

The Wiener process  $W = \{W(t) : t \in T^d\}$  on  $T^d$  is characterized by

(a)  $P\{W \in C_d\} = 1$ ;

(b) if  $B_1, B_2, \dots, B_k$  are pairwise disjoint blocks in  $T^d$ , then the incre-

ments  $W(B_1), W(B_2), \dots, W(B_k)$  are independent normal variables with means zero and variances  $|B_1|, |B_2|, \dots, |B_k|$ , where  $|\cdot|$  denotes the  $d$ -dimensional Lebesgue measure in  $T^d$ .

If  $\mathcal{A}$  is a class of sets, the Wiener process indexed in  $\mathcal{A}$ ,  $W = \{W(A): A \in \mathcal{A}\}$ , is the Gaussian centered process with covariances  $\text{cov}(W(A), W(B)) = |A \cap B|$  for  $A, B \in \mathcal{A}$ . If  $\mathcal{A}$  is the class of the blocks in  $T^d$ , denoted by  $\mathcal{B}^d$ ,  $\{W(B): B \in \mathcal{B}^d\}$  is the increments process of  $\{W(t): t \in T^d\}$  around the blocks  $B$ 's. It results that  $W = \{W(B): B \in \mathcal{B}^d\}$  is continuous [8].

### 3. MAIN THEOREM

Let  $T = [-a, a] \times [-b, b]$  be a half-open rectangle in  $R^2$ . We divide the plane in a grid  $\{T_j: j \in Z^2\}$ , where  $T_j$ 's are all rectangles with the same dimension as  $T$ , such that if  $j = (j_1, j_2)$ ,  $T_j$  has center  $C_j = (2aj_1, 2bj_2)$ . For  $\lambda = (\lambda_1, \lambda_2) \in Z^2$  we define the following dilation of  $T$

$$\lambda T = \begin{cases} T & \text{if } \lambda_1 = \lambda_2 = 1, \\ \bigcup_{j \in Q_\lambda} T_j & \text{if } \lambda_1 > 1 \text{ or } \lambda_2 > 1, \end{cases}$$

where  $Q_\lambda = \{j \in Z^2: -(\lambda-1) \leq j \leq \lambda-1\}$ . Note that  $|\lambda T| = |Q_\lambda| |T|$ , where  $|Q_\lambda| = \text{card}(Q_\lambda) = (2\lambda_1 - 1)(2\lambda_2 - 1)$  and

$$F(f, T) = |Q_\lambda|^{-1} \sum_{j \in Q_\lambda} F(f, T_j).$$

**THEOREM 1.** Suppose that the field  $X$  is strongly mixing and

- (i) the mixing satisfies  $\sum_{k=1}^{\infty} k^{(q+[q/2]) - 1} \alpha^{\delta/(q+\delta)}(k)$ ;
- (ii)  $E[F(f, T)]^{q+\delta} < \infty$  for some  $q \geq 2$  and  $\delta > 0$ ;
- (iii) the variance defined by Lemma 1 has the form

$$\sigma^2 = |T|^{-1} \int_{R^2} I(0, t, f) dt > 0,$$

where

$$I(s, t, f) = \int_{-\pi}^{\pi} f(\theta) d\theta \int_{-\pi}^{\pi} f(\varphi) H_{st}(x, \theta, \varphi) d\varphi$$

with

$$\begin{aligned} H_{st}(x, \theta, \varphi) &= P_{X(s), X(t), \Theta(s), \Theta(t)}(x, x, \theta, \varphi) \times \\ &\times E[|\dot{X}(s)| |\dot{X}(t)| | X(s) = X(t) = x, \Theta(s) = \theta, \Theta(t) = \varphi] - \\ &- P_{X(s), \Theta(s)}(x, \theta) E[|\dot{X}(s)| | X(s) = x, \Theta(s) = \theta] \times \\ &\times P_{X(t), \Theta(t)}(x, \varphi) E[|\dot{X}(t)| | X(t) = x, \Theta(t) = \varphi], \end{aligned}$$

where  $P_{X, \Theta}$  denotes the joint density of  $X$  and  $\Theta$ .

Then the field  $\{Z_\lambda(T): T \in \mathbb{R}^2\}$ , defined by

$$Z_\lambda(T) = \left[ \frac{|Q_\lambda| |T|}{\sigma^2} \right]^{1/2} [F(f, \lambda T) - EF(f, T)],$$

converges weakly to the Wiener process indexed in  $\mathcal{R}^2$ ,  $\{W(T): T \in \mathcal{R}^2\}$ .

Proof. We can define the following discrete random field: for every  $j \in \mathbb{Z}^2$ ,  $Y_j = F(f, T_j) - EF(f, T_j)$ . It is not difficult to see that if the field  $X$  is strongly mixing, then so is  $Y = \{Y_j: j \in \mathbb{Z}^2\}$ , and the mixing coefficient of  $Y$  is less than  $\alpha(k-h)$ , where  $\alpha$  is the strong mixing coefficient of  $X$  and  $h$  is the length of the diagonal of  $T$ . In fact, if  $N \in \mathbb{Z}^2$ , because  $F(f, T)$  is measurable we have

$$\sigma \{Y_j: j \in N\} \subset \sigma \{X(t): t \in \bigcup_{j \in N} T_j\}$$

and, therefore,

$$\sup \{ |P(A \cap B) - P(A)P(B)| : A \in \sigma(N), B \in \sigma(N') \} \leq \alpha \left( p \left( \bigcup_{j \in N} T_j, \bigcup_{j \in N'} T_j \right) \right).$$

We conclude that  $Y$  is strong mixing and its mixing coefficient satisfies

(i).

On the other hand,

$$Z_\lambda(T) = \frac{|T|^{1/2}}{(|Q_\lambda| \sigma^2)^{1/2}} \sum_{j \in Q_\lambda} Y_j = (|n| \sigma^2)^{-1/2} \sum_{-[nt] \leq j \leq [nt]} Y_j.$$

The last equation results from the change  $\lambda - 1 = [nt]$ . Accordingly, this suggests making use of a Functional Central Limit Theorem for mixing multiparametric processes.

PROPOSITION 1. Suppose  $Y = \{Y_j: j \in \mathbb{Z}^d\}$  is a strongly mixing centered stationary real random field such that, for some even number  $q \geq 2$  and  $\delta > 0$ ,

$$(i) \quad \sum_{k=1}^{\infty} k^{d(3/2)q-1} \alpha^{\delta/(q+\delta)}(k) < \infty,$$

$$(ii) \quad E|Y_0|^{q+\delta} < \infty.$$

If  $\sigma^2$ , defined by Lemma 1, is positive, then, for  $t \in T^d$ ,

$$Z_n(t) = (|n| \sigma^2)^{-1/2} \sum_{0 \leq j \leq [nt]} Y_j$$

converges weakly to the  $d$ -parameter Wiener process  $\{W(t): t \in T^d\}$ .

Proof. The technics used in this proof are essentially those of Billingsley [1] (Theorem 20.1) in the generalization made by Deo [5] for multiparametric processes. The conditions  $EZ_n(t) \rightarrow 0$  and  $EZ_n^2(t) \rightarrow |t|$  as  $n \rightarrow \infty$  are

trivially seen to be satisfied. It remains, therefore, to be proved the uniform integrability of  $Z_n^2(t)$  and its tightness.

In order to prove that we will make use of the following property: if the assumptions of Proposition 1 hold, then

$$(1) \quad E \left| \sum_{j \in A} Y_j \right|^q \leq K |A|^{q/2}$$

for some even integer, where  $K$  is a finite constant only depending on  $q$  and  $d$ , the moments of  $Y$  and  $\alpha$  [9].

The uniform integrability of  $Z_n^2(t)$  is a consequence of the previous property and the following inequality:

$$E(S_n^2 | |S_n^2| > a) \leq a^{(2-q)/2} |n|^{-q/2} E|S_n|^q.$$

The tightness condition comes from the following Oscillation Lemma, due to Doukhan-Portal [6]. This lemma is more powerful than the one we require because it gives a precise rate for oscillation, and we need to modify it a little, since originally it is referred for empirical process.

LEMMA 2. *With the same assumption of Proposition 1, for  $p \in N$  and some  $0 < \delta < 1$  such that  $p(1-\delta) > d$  we get*

$$P \left\{ \sup \{ |Z_n(s) - Z_n(t)| : \|s - t\| \leq n^{\beta-1} \} \geq kn^{-\theta} \right\} \leq kn^{-\theta},$$

where  $\theta = [p(1-\beta(1-\delta)) - d(1-\beta)] / (2p+1)$  and  $k$  is a constant depending only on  $p$  and  $d$ , the moments of  $Y$  and  $\alpha$ . Here  $\|s\| = \sum_{1 \leq j \leq d} |s_j|$ .

Proof. The proof comes from the inequality

$$E |S_{[nt]} - S_{[ns]}|^q \leq E \left| \sum_{[nt] \leq j \leq [ns]} Y_j \right|^q \leq K |n|^{q/2} \|t - s\|^{q/2}$$

that is a consequence of (1) and the proper definition of  $Z_n(t)$ . Lemma III.4 of Doukhan-Portal [6] holds trivially in this case and the rest of the proof follows without change.

The proof of Theorem 1 is completed by making use of Proposition 1 (taking  $d = 2$ ) and noting that  $Z_\lambda(T)$  can be expressed

$$Z_\lambda(T) = (|n| \sigma^2)^{-1/2} \left[ \sum_{0 \leq j \leq [nt]} Y_j + \sum_{[n(-t)] \leq j \leq 0} Y_j + \sum_{\substack{0 \leq j_1 \leq [n_1 t_1] \\ [n_2(-t_2)] \leq j_2 \leq 0}} Y_j + \sum_{\substack{[n_1(-t_1)] \leq j_1 \leq 0 \\ 0 \leq j_2 \leq [n_2 t_2]}} Y_j \right]$$

which converges weakly to

$$W(t_1, t_2) + W(-t_1, -t_2) - W(t_1, -t_2) - W(-t_1, t_2) = W(T).$$

#### 4. APPLICATIONS STUDY OF THE VARIANCE $\sigma^2$

The expression for the variance  $\sigma^2$  in Theorem 1 (iii) is a consequence of the Rice formulae for the moments of  $F(f, T)$ . Cabaña [3] has proved that if the field  $X$  has a Jacobian a.s. Lipschitz continuous and the Lipschitz

constant does not depend on  $\omega$ , for every  $t \in T$  the probability distribution of  $X(t)$  has density  $P_{X(t)}$ , there is a joint density of  $X$  and  $\dot{X}$  and it is continuous, and the function  $f$  is non-negative and Lipschitz continuous, then the Rice formulae hold even if they are infinite. Moreover, if  $X$  is Gaussian, instead of the first condition one can demand that  $\dot{X}$  and the derivative of its covariance function satisfy a Lipschitz condition in any compact  $K$ .

#### 4.1. Gaussian case.

LEMMA 3. Suppose  $X$  is a Gaussian field with covariance function  $\Gamma$ . Suppose also  $X$  satisfies the assumptions of the previous section. If  $\Gamma$  has two derivatives, and  $\Gamma(u)$ ,  $\dot{\Gamma}(u)$  and  $\ddot{\Gamma}(u)$  tend to zero as  $|u| \rightarrow \infty$ , then the variance has the form (iii) of Theorem 1.

Proof. Making use of Lemma 1, we have

$$|Q_\lambda| \text{Var} F(f, \lambda T) = |Q_\lambda|^{-1} \mathbb{E} \left| \sum_{j \in Q_\lambda} Y_j \right|^2 \rightarrow \sum_{j \in \mathbb{Z}^2} \mathbb{E}(Y_0 Y_j) = \sigma^2$$

as  $|\lambda| \rightarrow \infty$  and, therefore, we have to prove that

$$(2) \quad \left| |Q_\lambda| \text{Var} F(f, \lambda T) - \sigma^2 \right| \\ = \left| |Q_\lambda|^{-1} |\lambda T|^{-2} \int_{\lambda T \times (\lambda T)'} ds dt \int_{-\pi}^{\pi} f(\theta) d\theta \int_{-\pi}^{\pi} f(\varphi) H_{st}(x, \theta, \varphi) d\varphi \right|$$

tends to zero as  $\lambda \rightarrow \infty$ . Here  $(\lambda T)'$  is the complement of  $\lambda T$  in  $\mathbb{R}^2$ , and the equation follows from Theorem 2, the expression of the variance  $\sigma^2$  and stationarity of the field.

In order to prove that we will see that

$$(A) \quad \left| P_{X(s), X(t)}(x, x) - P_{X(s)}(x) P_{X(t)}(x) \right| \rightarrow 0 \quad \text{as } p(s, t) \rightarrow \infty$$

and

$$(B) \quad \left| P \{X(s) \in I, X(t) \in J / X(s) = X(t) = x\} - P \{X(s) \in I / X(s) = x\} \times \right. \\ \left. \times P \{X(t) \in J / X(t) = x\} \right| \rightarrow 0$$

as  $p(s, t) \rightarrow \infty$ , where  $I$  and  $J$  are both rectangles of  $\mathbb{R}^2$ . The proof will follow from (A) and (B) arguing as in the last part of the proof of Lemma 4.

We can assume without loss of generality that  $X$  is normalized in such a way that  $\Gamma(0) = -\dot{\Gamma}(0) = 1$ . We see that (B) is less than or equal to

$$\left| P \{ \dot{X}(s) \in I, \dot{X}(t) \in J / X(s) = X(t) = x \} - P \{ \dot{X}(s) \in I / X(s) = x \} \times \right. \\ \left. \times P \{ \dot{X}(t) \in J / X(s) = X(t) = x \} \right| + \\ + \left| P \{ \dot{X}(s) \in I / X(s) = X(t) = x \} - P \{ \dot{X}(s) \in I / X(s) = x \} \right| + \\ + \left| P \{ \dot{X}(t) \in J / X(s) = X(t) = x \} - P \{ \dot{X}(t) \in J / X(t) = x \} \right|.$$

Since  $X$  is Gaussian,  $X$  and  $\dot{X}$  conditioned by  $X$  are Gaussian as well, and we can write the first term as

$$(b) \quad |\mathbb{P}\{Z(s) \in I, Z(t) \in J\} - \mathbb{P}\{Z(s) \in I\} \mathbb{P}\{Z(t) \in J\}|,$$

where  $Z(s)$  and  $Z(t)$  are two Gaussian variables with parameters  $m_s = -m_t = \Gamma(u)x/(1+\Gamma(u))$ , where  $u = s-t$ ,

$$\begin{aligned} \text{Var } Z(s) &= \text{Var } Z(t) = 1 + \frac{\Gamma(u)\dot{\Gamma}(u)}{1-\Gamma^2(u)}, \\ \text{cov}(Z(s), Z(t)) &= -\dot{\Gamma}(u) + \frac{\Gamma(u)\dot{\Gamma}(u)\dot{\Gamma}(u)}{1-\Gamma^2(u)}. \end{aligned}$$

If  $\Sigma$  denotes the covariance matrix of the joint distribution of  $Z(s)$  and  $Z(t)$ , we denote by  $\Lambda$  the matrix which coincides with  $\Sigma$  in the diagonal and has zeros as other elements, and by  $m$  the column vector which has  $m_s$  and  $m_t$  as coordinates. Hence (b) can be written as

$$\begin{aligned} & \left| \frac{(\det \Sigma)^{-1/2}}{2\pi} \int_{I \times J} \exp\{-(y-m)\Sigma^{-1}(y-m)'\} dy - \right. \\ & \quad \left. - \frac{(\det \Lambda)^{-1/2}}{2\pi} \int_{I \times J} \exp\{-(y-m)\Lambda^{-1}(y-m)'\} dy \right| \\ & \leq \left| \frac{(\det \Sigma)^{-1/2} - (\det \Lambda)^{-1/2}}{2\pi} \right| \int_{I \times J} \exp\{-(y-m)\Lambda^{-1}(y-m)'\} dy + \\ & \quad + \frac{(\det \Sigma)^{-1/2}}{2\pi} \int_{I \times J} |1 - \exp\{-(y-m)[\Sigma^{-1} - \Lambda^{-1}](y-m)'\}| dy. \end{aligned}$$

Now according to the assumption and to the form of the variance and covariance, (b) tends to zero as  $|u| \rightarrow \infty$ .

The other terms have a similar behaviour, since both can be thought as  $|\mathbb{P}\{Z_1 \in I\} - \mathbb{P}\{Z_2 \in J\}|$ , where  $Z_1$  and  $Z_2$  are two Gaussian variables with parameters

$$m_1 = \frac{\dot{\Gamma}(u)x}{1+\Gamma(u)}, \quad m_2 = 0, \quad \text{Var } Z_1 = 1 + \frac{\dot{\Gamma}(u)\dot{\Gamma}(u)}{1-\Gamma^2(u)}, \quad \text{Var } Z_2 = 1.$$

Finally, (A) is deduced easily from the assumption, since the variables are centered, and the covariance matrix of  $(X(s), X(t))$  has the form

$$\begin{pmatrix} 1 & \Gamma(u) \\ \Gamma(u) & 1 \end{pmatrix}$$

which tends to the identity for  $|u| \rightarrow \infty$ .



#### 4.2. The \*-mixing case.

LEMMA 4. Suppose that  $X$  satisfies the Cabaña's assumptions (see the beginning of section 4) and also

(ii)  $P_{X(t)}(x)$  and  $E[|\dot{X}|/X = x]$  are uniformly bounded in a neighbourhood  $V$  of  $x$ ; then the variance  $\sigma^2$  has the form (iii) of Theorem 1.

Proof. Like Lemma 3 we have to prove that expression (2) tends to zero as  $|\lambda| \rightarrow \infty$ . This expression is smaller than

$$K^2 |Q_\lambda|^{-1} |T|^{-2} \int_{\lambda T \times (\lambda T)'} ds dt \int_{-\pi}^{\pi} \int_{-\pi}^{\pi} |H_{st}(x, \theta, \varphi)| d\theta d\varphi.$$

It is not difficult to see that

$$H_{st}(x, \theta, \varphi) = \lim_{\varepsilon \rightarrow 0} \varepsilon^{-4} \{E[Z_s(x, \theta, \varepsilon) Z_t(x, \varphi, \varepsilon)] - E Z_s(x, \theta, \varepsilon) E Z_t(x, \varphi, \varepsilon)\},$$

where

$$Z_s(x, \theta, \varepsilon) = \|\dot{X}(s)\| \mathbf{1}_{[x \leq \dot{X}(s) \leq x + \varepsilon]} \mathbf{1}_{|\theta \leq \theta(s) \leq \theta + \varepsilon|}.$$

Since the integrands are positive, an application of Fatou's Lemma and the inequality of moments for \*-mixing variables [7] give

$$\begin{aligned} & \int_{-\pi}^{\pi} \int_{-\pi}^{\pi} |H_{st}(x, \theta, \varphi)| d\theta d\varphi \\ & \leq \psi(p(s, t)) \liminf_{\varepsilon \rightarrow 0} \varepsilon^{-4} \int_{-\pi}^{\pi} E|Z_s(x, \theta, \varepsilon)| d\theta \int_{-\pi}^{\pi} E|Z_t(x, \varphi, \varepsilon)| d\varphi. \end{aligned}$$

Making use of the conditional expectation properties and Fubini's Theorem, we obtain that each integral of the right-hand member of the previous inequality is majorated by  $\varepsilon^2 CC'$ , where  $C$  and  $C'$  are bounds of  $P_{X(t)}(x)$  and  $E[|\dot{X}(s)|/X(s) = x]$  respectively, both uniform in a neighbourhood of  $x$  which contains the interval  $[x, x + \varepsilon]$ . Hence, substituting in the last expression, the limit in  $\varepsilon$  disappears and the proof will be completed if we show that

$$(c) \quad |Q_\lambda|^{-1} |T|^{-2} \int_{\lambda T \times (\lambda T)'} \psi(p(s, t)) ds dt \rightarrow 0 \quad \text{as } |\lambda| \rightarrow \infty.$$

Let us denote by  $m_\lambda$  and  $M_\lambda$  the sets

$$m_\lambda = \{(s, t) \in \mathbf{R}^4 : s \in \lambda T, t \notin \lambda T, p(s, t) \leq \mu(\lambda)\},$$

$$M_\lambda = \{(s, t) \in \mathbf{R}^4 : s \in \lambda T, t \notin \lambda T, p(s, t) > \mu(\lambda)\},$$

where  $\mu(\lambda) = o[(z\lambda_1 - 1)(z\lambda_2 - 1)]$  and  $\mu(\lambda) \rightarrow \infty$  as  $\lambda \rightarrow \infty$ . If  $L$  is a bound of  $\psi$ , then

$$\iint_{m_\lambda} \psi(p(s, t)) ds dt \leq L|m_\lambda| = L \times o[|Q_\lambda||T|].$$

On the other hand,

$$\iint_{n_\lambda} \psi(p(s, t)) ds dt = \int_{\lambda T} ds \int_{\{p(s, t) > \mu(\lambda)\}} \psi(p(s, t)) dt = |Q_\lambda||T| \int_{\{|u| > \mu(\lambda)\}} \psi(|u|) du,$$

therefore (c) is less than

$$|Q_\lambda|^{-1}|T|^{-2} L \times o[|Q_\lambda||T|] + |T|^{-1} \int_{\{|u| > \mu(\lambda)\}} \psi(|u|) du$$

which tends to zero as  $\lambda \rightarrow \infty$ .

### 5. EXISTENCE OF THE MOMENT OF $F(f, T)$

Assumption (ii) in Theorem 1 requires the finiteness of the  $q + \delta$  order moment. The Wschebor Theorem for the Rice formula [11] gives conditions for the validity of the formulae and also guarantees their finiteness. However the following lemma, suggested by Wschebor, shows in a geometrical way how the problem of moments existence can be studied without looking at the complicated integrals in the Rice formulae.

Let  $T = T_1 \times T_2$  be the Cartesian product of two intervals of  $R$ . We call 1-section of  $X$  determined by  $t_1$ , which will be denoted by  $X_{t_1}^1$ , the uniparametric process

$$X_{t_1}^1: T_2 \rightarrow R, \quad t \rightarrow X_{t_1}^1(t) = X(t_1, t).$$

In the same way we can define  $X_{t_2}^2$ . Obviously, the 1-sections are all measurable for every  $t_i$  ( $i = 1, 2$ ). Since  $X$  is stationary, it is sufficient to consider the  $i$ -sections  $X^i$  determined by  $t_i = 0$ . If  $X$  is a.s. of class  $C^p$ , then a.s. every  $i$ -section is a.s. of class  $C^p$ .

In what follows we will take  $f \equiv 1$  and write  $\mathcal{L}(\mathcal{C}_x) = F(1, T)$ . Let  $N_x^i(T_j)$  be the number of crossing of the process  $X^i$  with the level  $x$  in the interval  $T_j$ , where  $i \neq j$ ,  $i, j = 1, 2$ . Moreover, let

$$Z_p^i = \sup_{t_j \in T_j} |(X^i)^{(p)}(t_j)|.$$

LEMMA 5. Suppose that the field  $X$  and its derivative  $\dot{X}$  have a.s. continuous paths. Then

$$\mathcal{L}(\mathcal{C}_x) \leq \int_{T_1} N_x^1(T_2) dt_1 + \int_{T_2} N_x^2(T_1) dt_2.$$

COROLLARY 1. *With the same assumptions*

$$\{E[\mathcal{L}(\mathcal{C}_x)^p]\}^{1/p} \leq |T_1| \{E[N_x^1(T_1)]^p\}^{1/p} + |T_2| \{E[N_x^2(T_2)]^p\}^{1/p}.$$

COROLLARY 2. *Suppose that  $X$  is a Gaussian field, with covariance function  $\Gamma$ , which is normalized ( $\Gamma(0) = -\dot{\Gamma}(0) = 1$ ). If, for  $u, v \in \mathbf{R}$ ,*

$$\Gamma((0, u), (0, 0)) = E(X(0, u)X(0, 0)) = 1 + \frac{u^2}{2} + \frac{C|u|^3}{6} + o(u^3),$$

$$\Gamma((v, 0), (0, 0)) = E(X(v, 0)X(0, 0)) = 1 + \frac{v^2}{2} + \frac{C'|v|^3}{6} + o(v^3)$$

for  $u$  and  $v$  in a neighbourhood  $V$  of 0 and  $C$  and  $C'$  positive constants, then  $E[\mathcal{L}(\mathcal{C}_x)]^k < \infty$  for every  $k = 1, 2, \dots$

Proof. The proof follows from Lemma 5 and Theorem 4.1 in [4].

COROLLARY 3. *Suppose  $X$  has a.s. class  $C^p$  paths ( $p \geq 2$ ) and*

(i) *For every  $t_2 \in T_2$ ,  $(X^2(t_1), \dot{X}^2(t_1))$  has joint density uniformly bounded by  $C$  and for every  $t_1 \in T_1$ ,  $(X^1(t_2), \dot{X}^1(t_2))$  has joint density uniformly bounded by  $C$ .*

(ii) *Let  $m < 2p - 2$  if there is an  $r > 2m/(2p - 2 - m)$  such that  $E|Z_p^i|^r < \infty$  for every  $p = 1, 2, \dots$  and  $i = 1, 2$ .*

Then

$$\{E[\mathcal{L}(\mathcal{C}_x)]^m\}^{1/m} \leq L_1 |T_1| \{E|Z_p^2|^r\}^{1/m} + L_2 |T_2| \{E|Z_p^1|^r\}^{1/m} + L_3 < \infty,$$

where  $L_1, L_2$  and  $L_3$  depend only on  $p, m, r, |T|$  and  $C$ .

COROLLARY 4. *If  $X$  has a.s. class  $C^\infty$  paths, if (i) of Corollary 3 holds and  $E|Z_p^i| < \infty$  for every  $p = 1, 2, \dots$  and  $i = 1, 2$ , then*

$$E[\mathcal{L}(\mathcal{C}_x)]^m < \infty \quad \text{for every } m = 1, 2, \dots$$

Proofs of Corollaries 3 and 4 follow from Lemma 4 and Corollaries 3 and 4 of Wschebor [11] (p. 36).

Proof of Lemma 5. The proof is completely geometrical and the same idea can be extended for more general situations (dimension  $> 2$ ).

Suppose  $\mathcal{C}_x$  is a polygonal. Every segment of  $\mathcal{C}_x$  has length less than the sum of the lengths of its projections over the coordinates axes. If we take a partition in every interval  $T_1$  and  $T_2$ , the length of every semi-interval contributes to the sum every time the  $i$ -th section process determined by  $t_i$  in the semi-interval, crosses the level  $x$  in the interval  $T_j$  ( $i \neq j$ ). Therefore

$$\mathcal{L}(\mathcal{C}_x) \leq \sum_{k=1}^N N_x^1(T_2) [t_1^k - t_1^{k-1}] + \sum_{h=1}^M N_x^2(T_1) [t_2^h - t_2^{h-1}],$$

where  $\{t_1^0, t_1^1, \dots, t_1^N\}$  and  $\{t_2^0, t_2^1, \dots, t_2^M\}$  are both partitions of  $T_1$  and  $T_2$ , respectively. Lemma 5 follows by taking the limit when the sizes of the partitions tend to zero.

**Acknowledgement.** I am greatly indebted to my thesis adviser, Enrique Cabaña, who formulated this problem and suggested many crucial ideas. Also, I would like to thank J. R. León, M. Wschebor and J. Ortega for many helpful discussions. The improvement in section 5 of Wschebor's results were suggested by him. I thank P. Doukhan who gave me invaluable advise in the organization of this paper.

#### REFERENCES

- [1] P. Billingsley, *Convergence of probability measures*, Wiley, New York 1969.
- [2] E. Cabaña, *Affine processes, a test of isotropy* (to appear in SIAM (1984)).
- [3] — *Esperanza para integrales sobre conjuntos de nivel aleatorios*, Actas del II Congreso Latinoamericano de Probabilidades y Estadística Matemática, 1985, p. 65–82.
- [4] J. Cuzick, *Conditions for finite moments of the number of zero crossing for Gaussian processes*, Ann. Probab. 5.3 (1975), p. 849–858.
- [5] Ch. M. Deo, *A functional central limit theorem for stationary random fields*, ibidem 4.3 (1975), p. 708–715.
- [6] P. Doukhan et F. Portal, *Principe d'invariance faible pour la fonction de répartition empirique dans un cadre multidimensionnel et mélangeant* (to appear).
- [7] P. Doukhan et J. R. León, *Quelques notions de mélanges et des exemples de processus mélangeants*, Prépublication 85 T20, Université Paris-Sud, Orsay 1985.
- [8] R. M. Dudley, *Sample functions of the Gaussian processes*, No. 1, Vol. 1 (1973), p. 66–103.
- [9] X. Guyon et S. Richardson, *Vitesse de convergence du théorème de la limite centrale pour des champs faiblement dépendants*, Z. Wahrsch. verw. Gebiete 66 (1984), p. 297–314.
- [10] M. Wschebor, *Formule de Rice en dimension  $d$* , ibidem 60 (1982), p. 393–401.
- [11] — *Surfaces aléatoires. Mesures géométriques des ensembles de niveau*, Lecture Notes in Math. 1147 (1985), Springer.

Departamento de Matemáticas  
Instituto Venezolano de Investigaciones Científicas (IVIC)  
Apartado 21201  
Caracas 1020-A  
Venezuela

Received on 24. 9. 1987