

SKOROKHOD PROBLEM - ELEMENTARY PROOF OF THE AZEMA-YOR FORMULA

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Abstract. Let μ be a centered probability measure with the finite second moment. Let the stopping time T for the Brownian motion W be defined as

$$T = \inf \{t \geq 0; \Psi(W_t) \leq \sup_{0 \leq s \leq t} W_s\},$$

where Ψ is a barycenter function of measure μ . Azema and Yor [1] have shown that W_T has then the distribution μ and $ET = \int x^2 \mu(dx)$. This paper contains an elementary proof of this result.

Introduction. Skorokhod [8] has shown that for the centered probability measure μ with a continuous distribution function there exists the Brownian motion W and the stopping time T so that the distribution of W_T is μ . Moreover, if μ has the finite second moment, then

$$ET = \int_{-\infty}^{+\infty} x^2 \mu(dx).$$

That construction was improved by Monroe [5]. However, the stopping time T , given by the Skorokhod's construction is regarded with respect to the filtration essentially bigger than the natural filtration of Brownian motion. Dubins [4], Rost [7], Chacon-Walsh [3] and Azema-Yor [1] gave new constructions of stopping times with the desired property but which are stopping times with respect to the natural filtration of Brownian motion. Construction given by Azema-Yor is the best one in some respects. It is an explicit formula and not a result of a limit procedure. Pierre [6] gave a new proof of the Azema-Yor formula but with assumptions of regularity of a

measure μ . The proof of this formula given in our paper is based on the following

PROPERTY. Let W be the Brownian motion. If $W_0 \equiv 0$, $a < 0 < b$,

$$\tau_x = \inf\{t \geq 0: W_t = x\}, \quad x \in \mathbb{R},$$

then

$$P(\tau_a < \tau_b) = \frac{b}{b-a}.$$

Definition 1. If μ is a probability distribution such that

$$\int_{-\infty}^{\infty} |x| \mu(dx) < \infty,$$

then the *barycenter function* Ψ of the measure μ is defined as

$$\Psi(a) = \begin{cases} \frac{1}{\mu([a, \infty))} \int_{[a, \infty)} x \mu(dx) & \text{if } \mu([a, \infty)) > 0, \\ a & \text{otherwise.} \end{cases}$$

Notation. We write

$$S_t = \sup_{0 \leq s \leq t} W_s.$$

THEOREM. If $T = \inf\{t \geq 0; \Psi(W_t) \leq S_t\}$, then W_T has a distribution μ . Moreover, if

$$\int_{-\infty}^{+\infty} x^2 \mu(dx) < \infty,$$

then

$$ET = \int_{-\infty}^{+\infty} x^2 \mu(dx).$$

Remark 1. T is a stopping time with respect to the natural filtration of Brownian motion.

Remark 2. To show that W_T has a distribution μ it is enough to prove the following implication:

$$\mu((a, \infty)) > 0 \Rightarrow \mu([a, \infty)) = P(\{W_T \geq a\}).$$

Indeed, if $\mu((a, \infty)) = 0$ and $a_n \nearrow a$ ($n = 1, 2, \dots$), then either there exists an N such that $\mu([a_n, \infty)) = 0$ for $n \geq N$ and then, for $b \geq 0$, $\Psi(a_N + b) =$

$a_N + b$ so that $W_T \leq a_N$ and $P(\{W_T \geq a\}) = 0$ or, for every $n = 1, 2, \dots$, $\mu((a_n, \infty)) > 0$ and then

$$\mu([a, \infty)) = \lim_{n \rightarrow \infty} \mu([a_n, \infty)) = \lim_{n \rightarrow \infty} P(\{W_T \geq a_n\}) = P(\{W_T \geq a\}).$$

Notation. We write

$$P_a = P(\{W_T \geq a\}), \quad \varphi(a) = \inf_{x \leq a} \{\Psi(x) - x\}, \quad v(x) = \inf \{y: \Psi(y) > x\},$$

$$K_i^n = i 2^{-n} \Psi(a) \quad (0 \leq i \leq 2^n), \quad K = K_{2^n}^n,$$

$$a_i^n = v(K_i^n) \quad (0 \leq i \leq 2^n), \quad a_0 = a_0^n, \quad \frac{-\infty}{-\infty} = 1.$$

Remark 3. Since the set where function Ψ is not continuous is at least denumerable, we can make an assumption that a is such that $\Psi(a_i^n) = K_i^n$ for $0 \leq i \leq 2^n$.

Remark 4. If $\Psi(a) > a$, then $\varphi(a) > 0$.

Proof. Suppose that $\varphi(a) = 0$. Since $\varphi(a) = \inf \{\Psi(x) - x; \varphi(a) - a \leq x \leq a\}$, there exists an $(x_n)_{n=0}^\infty \in [\varphi(a) - a, a]$ such that

$$\varphi(a) = \lim_{n \rightarrow \infty} (\Psi(x_n) - x_n) \quad \text{and} \quad x_n \rightarrow x_0.$$

If we can find a subsequence $x_{n_k} \nearrow x_0$, then — by the left continuity of the function φ — we have $\Psi(x_0) = x_0$. Also, if there exist subsequences $x_{n_k} \searrow x_0$, then

$$x_0 \leq \Psi(x_0) \leq \lim_{k \rightarrow \infty} \Psi(x_{n_k}) = x_0.$$

Since $x_0 \leq a$, we obtain a contradiction: $\Psi(a) \leq a$.

Remark 5. Let

$$\Gamma = \bigcup_{y \in R} y \times [\Psi(y), \lim_{z \searrow y} \Psi(z)]$$

and let $Z_t = (W_t, S_t)$ be the process with values in R^2 . Then, by the definition of the stopping time T , we infer that T is a first entrance time of process Z to the closed set Γ .

LEMMA 1. If $v(x) \leq y \leq x$, then

$$\frac{y - v(y)}{x - v(y)} \geq P(S_T \geq x | S_T \geq y) \geq \frac{y - v(x)}{x - v(x)}.$$

Proof. We have

$$\begin{aligned}
 & P(S_T \geq x | S_T \geq y) \\
 &= P(\{\text{after exit from } (y, y) \text{ } Z \text{ achieves } (x, x) \text{ before it enters } \Gamma\}), \\
 & \frac{y-v(y)}{x-v(y)} \\
 &= P(\{\text{after exit from } (y, y) \text{ } Z \text{ achieves } (x, x) \text{ before it enters } \{v(y) \times R\}\}), \\
 & \frac{y-v(x)}{x-v(x)} \\
 &= P(\{\text{after exit from } (y, y) \text{ } Z \text{ achieves } (x, x) \text{ before it enters } \{v(x) \times R\}\}).
 \end{aligned}$$

LEMMA 2. If $\Psi(a) > a$, then

$$\lim_{n \rightarrow \infty} \prod_{i=1}^{2^n} \frac{K_{i-1}^n - a_{i-1}^n}{K_i^n - a_i^n} \geq P_a \geq \lim_{n \rightarrow \infty} \prod_{i=1}^{2^n} \frac{K_{i-1}^n - a_i^n}{K_i^n - a_i^n}.$$

Proof. Let n be so large that $2^{-n} \Psi(a) < \varphi(a)$. From the definition, $\{W_T \geq a\} = \{S_T \geq \Psi(a)\}$, whence

$$P(\{W_T \geq a\}) = P(\{S_T \geq \Psi(a)\}) = \prod_{i=1}^{2^n} P(S_T \geq K_i^n | S_T \geq K_{i-1}^n).$$

Since

$$K_i^n - a_i^n \geq \Psi(a_i^n) - a_i^n \geq \varphi(a) > 2^{-n} \Psi(a) = K_i^n - K_{i-1}^n \quad (0 < i \leq 2^n),$$

we have $a_i^n < K_{i-1}^n < K_i^n$ and from Lemma 1 it follows that

$$\prod_{i=1}^{2^n} \frac{K_{i-1}^n - a_{i-1}^n}{K_i^n - a_{i-1}^n} \geq P_a \geq \prod_{i=1}^{2^n} \frac{K_{i-1}^n - a_i^n}{K_i^n - a_i^n}.$$

LEMMA 3. We have

$$\lim_{n \rightarrow \infty} \prod_{i=1}^{2^n} \frac{K_{i-1}^n - a_{i-1}^n}{K_i^n - a_{i-1}^n} = \lim_{n \rightarrow \infty} \prod_{i=1}^{2^n} \frac{K_{i-1}^n - a_i^n}{K_i^n - a_i^n}.$$

Proof. We can write

$$\begin{aligned}
 1 & \geq \left(\prod_{i=1}^{2^n} \frac{K_{i-1}^n - a_i^n}{K_i^n - a_i^n} \right) \left(\prod_{i=1}^{2^n} \frac{K_{i-1}^n - a_{i-1}^n}{K_i^n - a_{i-1}^n} \right)^{-1} \\
 &= \prod_{i=1}^{2^n-1} \frac{(K_{i-1}^n - a_i^n)(K_{i+1}^n - a_i^n)}{(K_i^n - a_i^n)^2} \frac{K_{2^n-1}^n - a_{2^n}^n}{K_{2^n}^n - a_{2^n}^n} \frac{K_1^n - a_0^n}{K_0^n - a_0^n} \\
 &= \prod_{i=1}^{2^n-1} \frac{(K_i^n - a_i^n - 2^{-n} \Psi(a))(K_i^n - a_i^n + 2^{-n} \Psi(a))}{(K_i^n - a_i^n)^2} \times
 \end{aligned}$$

$$\begin{aligned} & \times \frac{\Psi(a) - v(\Psi(a)) - 2^{-n} \Psi(a)}{\Psi(a) - v(\Psi(a))} \frac{2^{-n} \Psi(a) - v(0)}{-v(0)} \\ & = \prod_{i=1}^{2^n-1} \left[1 - \left(\frac{\Psi(a)}{2^n(K_i^n - a_i^n)} \right)^2 \right] \frac{\Psi(a) - v(\Psi(a)) - 2^{-n} \Psi(a)}{\Psi(a) - v(\Psi(a))} \frac{2^{-n} \Psi(a) - v(0)}{-v(0)} \\ & \geq \left[1 - \left(\frac{\Psi(a)}{2^n \varphi(a)} \right)^2 \right]^{2^n-1} \frac{\Psi(a) - v(\Psi(a)) - 2^{-n} \Psi(a)}{\Psi(a) - v(\Psi(a))} \frac{2^{-n} \Psi(a) - v(0)}{-v(0)} \rightarrow 1, \end{aligned}$$

when $n \rightarrow \infty$.

LEMMA 4. If $2^{-n} \Psi(a) < \varphi(a)$, then

$$\prod_{i=1}^{2^n} \frac{K_{i-1}^n - a_{i-1}^n}{K_i^n - a_i^n} \geq \mu([a, \infty)) \geq \prod_{i=1}^{2^n} \frac{K_{i-1}^n - a_i^n}{K_i^n - a_i^n}.$$

Proof. Since

$$\mu([a, \infty)) = \prod_{i=1}^{2^n} \frac{\mu([a_i^n, \infty))}{\mu([a_{i-1}^n, \infty))},$$

it is enough to show that

$$\frac{K_{i-1}^n - a_{i-1}^n}{K_i^n - a_{i-1}^n} \geq \frac{\mu([a_i^n, \infty))}{\mu([a_{i-1}^n, \infty))} \geq \frac{K_{i-1}^n - a_i^n}{K_i^n - a_i^n} \quad \text{for } 0 < i \leq 2^n.$$

From the definition of function Ψ and from Remark 3 we get

$$\frac{K_{i-1}^n - a_{i-1}^n}{K_i^n - a_{i-1}^n} = \frac{\mu([a_i^n, \infty))}{\mu([a_{i-1}^n, \infty))} \frac{\int_{[a_{i-1}^n, \infty)} (x - a_{i-1}^n) \mu(dx)}{\int_{[a_i^n, \infty)} (x - a_{i-1}^n) \mu(dx)} \geq \frac{\mu([a_i^n, \infty))}{\mu([a_{i-1}^n, \infty))}$$

and

$$\frac{K_{i-1}^n - a_i^n}{K_i^n - a_i^n} = \frac{\mu([a_i^n, \infty))}{\mu([a_{i-1}^n, \infty))} \frac{\int_{[a_{i-1}^n, \infty)} (x - a_i^n) \mu(dx)}{\int_{[a_i^n, \infty)} (x - a_i^n) \mu(dx)} \leq \frac{\mu([a_i^n, \infty))}{\mu([a_{i-1}^n, \infty))}.$$

Lemmas 2, 3, and 4 now easily imply that W_T has the distribution μ .

Remark 6. If T is a stopping time such that $ET < \infty$, then $EW_T = 0$ and $EW_T^2 = ET$.

To complete the proof of the theorem it must be shown that if

$$\int_{-\infty}^{+\infty} x^2 \mu(dx) < \infty,$$

then

$$ET = \int_{-\infty}^{+\infty} x^2 \mu(dx).$$

From Remark 6 it follows that it is enough to show that $ET < \infty$. Let Ψ_n and T_n ($n = 1, 2, \dots$) be defined as follows:

$$\Psi_n(x) = \begin{cases} 0 & \text{if } x \leq -n, \\ \Psi(x) & \text{if } -n < x \leq n, \\ \Psi(n) & \text{if } n < x \leq \Psi(n), \\ x & \text{if } \Psi(n) < x, \end{cases} \quad T_n = \inf \{t \geq 0; \Psi_n(W_t) \leq S_t\}.$$

From the definition, $T_n \rightarrow T$ and $ET_n < \infty$ ($ET_n \leq$ means exit time of Brownian motion from $[-n, \Psi(x)]$). So $ET_n = EW_{T_n}^2$.

To obtain from Fatou Lemma that $ET < \infty$ it is enough to show that

$$\limsup_{n \rightarrow \infty} EW_{T_n}^2 < \infty.$$

Let $A = \{W_T \in [n, \Psi(n)]\}$. Then

$$\begin{aligned} EW_{T_n}^2 &= E_{x_{Ac}} W_{T_n}^2 + E_{x_A} W_{T_n}^2 \leq E_{x_{Ac}} W_T^2 + E_{x_A} W_{T_n}^2 \\ &= EW_T^2 + E_{x_A} (W_{T_n}^2 - W_T^2) \leq EW_T^2 + 2\Psi(n) E_{x_A} (W_{T_n} - W_T) \\ &= EW_T^2 + 2\Psi(n) \int_{[n, \Psi(n)]} (\Psi(n) - x) \mu(dx) \\ &= EW_T^2 + 2\Psi(n) \int_{[\Psi(n), \infty)} (x - \Psi(n)) \mu(dx) \\ &\leq EW_T^2 + 2 \int_{[\Psi(n), \infty)} x^2 \mu(dx) \leq EW_T^2 + \varepsilon \quad \text{for } n \geq n_0, \end{aligned}$$

which completes the proof.

REFERENCES

- [1] J. Azema and M. Yor, *Une solution simple au problème de Skorokhod*, Séminaire de Probabilités XIII, Université de Strasbourg, 1977/78, Lecture Notes in Math. 721, p. 90-115, Springer-Verlag, New York 1979.
- [2] — *Le problème de Skorokhod: compléments à l'exposé précédent*, ibidem, p. 625-633.
- [3] R. Chacon and J. B. Walsh, *One-dimensional potential embedding*, Séminaire de Probabilités X, Université de Strasbourg, 1974/75, Lecture Notes in Math. 511, p. 19-23, Springer-Verlag, New York 1976.
- [4] L. E. Dubins *On a theorem of Skorokhod*, Ann. Math. Statist. 39 (1968), p. 2094-2097.
- [5] J. Monroe, *Processes that can be embedded in Brownian motion*, Ann. Probab. 6.1. (1978), p. 42-56.
- [6] M. Pierre, *Le problème de Skorokhod: une remarque sur la démonstration d'Azema-Yor*,

- Séminaire de Probabilités XIV (1978/79), Lecture Notes in Math. 784, p. 392-396, Springer-Verlag, New York 1980.
- [7] H. Root, *The existence of certain stopping times on Brownian motion*, Ann. Math. Statist. 40 (1969), p. 715-718.
- [8] A. W. Skorokhod, *Studies in the theory of random processes*, Addison-Wesley, Reading 1965.

SUPPLEMENTARY REFERENCES

- [A] D. Kennedy, *Some martingales related to cumulative sum tests and single-server queues*, Stochastic Processes and their Applic. 4 (1976), p. 261-269.
- [B] J. P. Lechoczky, *Formulas for stopped diffusion processes with stopping times based on the maximum*, Ann. Probab. 5.4 (1977), p. 601-607.

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Received on 12. 10. 1981

