## ON AN INVARIANCE PRINCIPLE FOR UNIFORMLY STRONG MIXING STATIONARY SEQUENCES WHEN $\mathcal{E}X^2 = \infty$

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Abstract. We prove that for uniformly strong mixing strictly stationary sequences a weak invariance principle holds for random variables with the second moment divergent. This is an extension of the result of Peligrad [8] for random variables with finite variance.

1. Introduction and notation. Let  $\{X_k\}_{k\in\mathbb{Z}}$  be a strictly stationary random sequence on probability space  $(\Omega, \mathcal{F}, \mathcal{P})$  and let  $\mathcal{F}_k^m$  denote the  $\sigma$ -field generated by  $\{X_i; m \leq i \leq k\}$ . Define:

$$\begin{split} \varphi_{\mathbf{n}} &= \varphi_{\mathbf{n}}(\{X_k\}) = \sup\{|\mathscr{P}(B/A) - \mathscr{P}(B)|; \ A \in \mathscr{F}_{-\infty}^0 \ , \ B \in \mathscr{F}_{\mathbf{n}}^\infty \ , \ \mathscr{P}(A) > 0\}, \\ \varrho_{\mathbf{n}} &= \varrho_{\mathbf{n}}(\{X_k\}) = \sup\{|\mathrm{Corr}(f,\,g)|; \ f, \ g - \mathrm{real}, \ f \in L^2(\mathscr{F}_{-\infty}^0), \ g \in L^2(\mathscr{F}_{\mathbf{n}}^\infty)\}. \end{split}$$

The sequence  $\{X_k\}_k$  is said to be uniformly strong mixing or  $\varphi$ -mixing if  $\lim_{n\to\infty} \varphi_n = 0$ . It is well known that  $\varrho_n \leq 2\varphi_n^{1/2}$ .

In this note, unless otherwise stated, we shall deal with strictly stationary  $\varphi$ -mixing sequences only.

Let  $S_n = \sum_{k=1}^n X_k$  and define the random element in  $\mathcal{D}(0, 1]$ :

$$\mathscr{X}_n(t) = \sigma_n^{-1} S_{[nt]}, \quad t \in (0, 1],$$

where  $\sigma_n^2 = \operatorname{Var} S_n$  and [ ] denotes the greatest integer function.  $\mathscr{X}_n$  satisfies the weak invariance principle (WIP) if  $\mathscr{X}_n$  converges weakly  $(\Rightarrow_w)$  to the standard Wiener measure  $\mathscr{W}$ .

Peligrad [8] proved that in the case  $\mathscr{E}X_1^2 < \infty$  WIP is equivalent to the Lindenberg condition. On the other hand, in the iid case the Central Limit Theorem holds for random variables with the second moment barely divergent [2].

The purpose of this note is to formulate and prove a WIP when  $\mathscr{E}X_1^2 = \infty$ . We use the following notation: let  $b_n \to_n + \infty$  for every  $n \in \mathbb{N}$  and denote by

 $\{\hat{X}_k\}_k$  an independent copy of  $\{X_k\}_k$ ;

$$\begin{split} X_{i}^{n} &= X_{i}I(|X_{i}| < b_{n}) - \mathscr{E}X_{i}I(|X_{i}| < b_{n}); \\ \hat{X}_{i}^{n} &= \hat{X}_{i}I(|\hat{X}_{i}| < b_{n}) - \mathscr{E}\hat{X}_{i}I(|\hat{X}_{i}| < b_{n}); \\ U_{i}^{n} &= X_{i}^{n} - \hat{X}_{i}^{n}; \quad T_{k}^{n} &= \sum_{i=1}^{k} X_{i}^{n}; \quad Z_{k}^{n} &= \sum_{i=1}^{k} U_{i}^{n}; \quad T_{n} &= T_{n}^{n}; \quad Z_{n} &= Z_{n}^{n}; \\ Y_{i}^{n} &= X_{i}I(|X_{i}| \geqslant b_{n}); \quad R_{k}^{n} &= \sum_{i=1}^{k} Y_{i}^{n}; \quad R_{n} &= R_{n}^{n}; \\ \hat{S}_{n} &= \sum_{i=1}^{n} \hat{X}_{i}; \quad (\tau_{k}^{n})^{2} &= \text{Var } T_{k}^{n}; \quad (z_{k}^{n})^{2} &= \text{Var } Z_{k}^{n}; \quad \tau_{n} &= \tau_{n}^{n}; \\ Z_{n} &= Z_{n}^{n}; \quad \mathscr{W}_{n}'(t) &= \tau_{n}^{-1} T_{n}^{n}; \quad \mathscr{W}_{n}''(t) &= \tau_{n}^{-1} S_{n}; \\ \mathscr{W}_{n}(t) &= \tau_{n}^{-1} \big( S_{n} - [nt] \mathscr{E}X_{1} I(|X_{1}| < b_{n}) \big). \end{split}$$

The Theorem we shall prove, in the case  $b_n = +\infty$  for all  $n \in \mathbb{N}$ , is Corollary 2.2 in [8]. As an application two corollaries will be proved, the second one is a recent result of Peligrad [9].

- 2. Auxiliary results and definitions. In this section we group some facts adapted for this note from more general theorems.
- (2.1)  $\{\max_{1 \le i \le n} \tau_n^{-2}(X_i^n)^2\}_n$  is uniformly integrable if and only if so is  $\{\max_{1 \le i \le n} \tau_n^{-2}(T_i^n)^2\}_n$

(see the proof of Proposition 2.1 in [8]).

(2.2) Let  $\{X_k\}_k$  be a centered  $L^2$ -stationary random sequence; then

$$(1 - \varrho_p)^{1/2} \max_{1 \le i \le n} \sigma_i \le \sigma_n + 2p\sigma_1$$

(see Lemma 4.2 in [7]).

(2.3) For any  $\{X_k\}_k$  such that

$$\varphi_1 + \max_{1 \leq i \leq n} \mathscr{P}(|S_n - S_i| > x_0) \leq \eta < 1,$$

for  $x \ge x_0$  we have

$$\mathscr{P}(\max_{1 \leq i \leq n} |S_i| > 2x) \leq (1 - \eta)^{-1} \mathscr{P}(|S_n| > x)$$

(see Lemma 1.1.6 in [4]).

(2.4) Let  $\{X_k^*\}_k$  denote an iid sequence with  $\mathcal{L}(X_1^*) = \mathcal{L}(X_1)$ ; then for x > 0:

$$(1-\varphi_1)\mathscr{P}(\max_{1\leqslant i\leqslant n}|X_i^*|>x)\leqslant \mathscr{P}(\max_{1\leqslant i\leqslant n}|X_i|>x)\leqslant (1+\varphi_1)\mathscr{P}(\max_{1\leqslant i\leqslant n}|X_i^*|>x)$$

(see Proposition 3.1 in [9]).

(2.5)  $\mathcal{L}(X_1)$  is said to be in the domain of attraction of the normal law  $(\mathcal{L}(X_1) \in \mathcal{D} \mathcal{A}(2))$  if there exist sequences  $\{A_n\}_n$  and  $\{b_n\}_n$  such that

$$\mathscr{L}(b_n^{-1}\sum_{k=1}^n X_i^* - A_n) \xrightarrow{\mathbf{w}} \mathscr{N}(0, 1), \quad n \to +\infty.$$

This is equivalent [2] to the slow variation of  $\mathscr{E}X_1^2I(|X_1| < x)$ , and then  $b_n := \inf\{x; \ x^{-2}\mathscr{E}X_1^2I(|X_1| < x) \le 1/n\}.$ 

(2.6) If  $\mathscr{E}X_1^2I(|X_1| < x)$  is slowly varying, then for  $\{b_n\}_n$  from (2.5) we obtain

$$\frac{n}{b_n}\mathscr{E}|X_1|I(|X_1|>b_n)\stackrel{n}{\longrightarrow}0, \quad n\to+\infty$$

(this follows easily from Theorem 2, VIII, §9, in [2]).

- (2.7) If  $x^2 \mathcal{P}(|X_1| > x)$  is a slowly varying function, then so is  $\mathcal{E}X_1^2 I(|X_1| < x)$  (see the same Theorem as in (2.6)); however, according to Exercise 32, VII, § 10, in [2], the converse is not true.
- (2.8) If  $x^2 \mathcal{P}(|X_1| > x)$  is a slowly varying function, then  $n\mathcal{P}(|X_1| > a_n) \xrightarrow{n} 1, \quad a_n = \inf\{x; \mathcal{P}(|X_1| > x) \le 1/n\}$

(see Lemma 1.8 in [10]).

(2.9) If  $x^2 \mathcal{P}(|X_1| > x)$  is a slowly varying function, then  $\mathcal{E}|X_1|I(|X_1| > x) \sim 2x\mathcal{P}(|X_1| > x), \quad x \to +\infty$ 

(see Theorem 8.1.4 in [1]).

(2.10) Assume  $n\mathcal{P}(|X_1| > b_n) \xrightarrow{n} 0$ , and  $\tau_n \to +\infty$ ,  $n \to +\infty$ , and  $\{\tau_n^{-2} T_{nj_n}^{2n}\}$  is uniformly integrable. Then

$$(W'_n(1)) \xrightarrow{\mathbf{w}} \mathcal{N}(0, 1), \quad n \to +\infty$$

(see Theorem 3 in [6]).

## 3. Results and proofs.

THEOREM. Assume that

(3.1) 
$$\lim n\mathscr{P}(|X_1| > b_n) = 0,$$

$$\lim_{n\to\infty}\tau_n=+\infty,$$

(3.3) 
$$\lim_{n\to\infty} \tau_n^{-2} \mathscr{E}\left(\max_{1\leqslant i\leqslant n} (X_i^n)^2\right) = 0.$$

Then

$$\mathscr{W}_n \xrightarrow{w} \mathscr{W}, \quad n \to +\infty.$$

Conversely, if  $\varphi_1 < 1$  and (3.4) holds, then (3.3) is satisfied.

COROLLARY 1. Let  $\mathcal{L}(X_1) \in \mathcal{DA}(2)$ ,  $\mathcal{E}X_1 = 0$  and

(3.5) 
$$\liminf_{n\to\infty} \tau_n b_n^{-1} > 0,$$

where  $b_n$  is defined in (2.5). Then

$$\mathcal{W}_n'' \stackrel{\mathsf{w}}{\Longrightarrow} \mathcal{W}, \quad n \to +\infty.$$

COROLLARY 2. Assume  $x^2 \mathcal{P}(|X_i| > x)$  is slowly varying,  $\mathcal{E}X_i = 0$ ,  $\varphi_1 < 1$ . Then (3.6) holds, and

(3.7) 
$$\sqrt{\pi/2} \mathscr{E}|S_n| \sim \tau_n, \quad n \to +\infty,$$

for some  $\{b_n\}_n$ .

Proof of the Theorem. We shall consider only the case  $\mathscr{E}X_1^2 = \infty$ , i.e.,  $b_n \stackrel{n}{\longrightarrow} +\infty$ , since the other case can be proved analogously. From (3.1) we see that

$$\max_{1 \le k \le n} \tau_n^{-1} |R_k^n| \xrightarrow{\mathscr{P}} 0, \quad n \to +\infty.$$

Thus in the proof we can restrict ourselves to  $W'_n$  random elements.

The direct half. An examination of the proof of Theorems 1 and 2 in [5] shows that it is enough to prove that

$$\max_{1 \le i \le [n\delta_{-1}]} \frac{(\tau_i^n)^2}{(\tau_n)^2} \xrightarrow{n} 0, \quad n \to +\infty,$$

for any  $\{\delta_n\}_n$  such that  $\lim_n \delta_n = 0$ . By (2.2), for any  $\varepsilon > 0$  and  $n \in \mathbb{N}$  such that  $\delta_n \leq \varepsilon$ , we have

$$\max_{1 \leq i \leq \lceil n\delta_n \rceil} \frac{\tau_i^n}{\tau_n} \leq (1 - \varrho_p)^{-1/2} \left( \frac{\tau_{\lceil n\epsilon \rceil}^n}{\tau_n} + 2p \frac{\tau_1^n}{\tau_n} \right),$$

so the required condition is satisfied if  $(\tau_n)^2$  is a regularly varying sequence with index 1 (see [1], p. 52), and

(3.8) 
$$\frac{\left(\tau_{[nt]}^n\right)^2}{\left(\tau_{[nt]}\right)^2} \xrightarrow{n} 1, \quad t \in (0, 1], \ n \to +\infty.$$

From (2.1) we infer that  $\{\tau_n^{-2} T_n^2\}_n$  is uniformly integrable, so by (2.10) and (3.1) we obtain

$$\mathscr{L}(z_{[nt]}^{-1}Z_{[nt]}^n) \xrightarrow{\mathbf{w}} \mathscr{N}(0,1), \quad n \to +\infty.$$

On the other hand, by (2.2) we have

$$(\tau_n)^2 = \mathscr{E}\left(\sum_{j=1}^{[n/[ht]]} \sum_{i=1}^{[nt]} X_{[nt](j-1)+i}^n + \sum_{i=[n/[nt]][nt]+1}^n X_i^n\right)^2$$

$$\leq 2^{n/[nt]} (\tau_{[nt]}^n)^2 + 2 \max_{1 \leq k \leq [nt]} (\tau_k^n)^2$$

$$\leq (\tau_{[nt]}^n)^2 (2^{2/t} + 4(1-\rho_n)^{-1}) + 8p^2 (1-\rho_n)^{-1} (\tau_1^n)^2.$$

so there exists a constant  $C = C(\varrho_p, t)$  such that

(3.10) 
$$\liminf_{n\to\infty} \tau_n^{-1} \tau_{[nt]}^n \geqslant C > 0,$$

since  $\lim_{n\to\infty} \tau_n^{-1} \tau_1^n = 0$  by (3.3). From (3.10) and (2.1) we infer that  $\{(\tau_{[nt]}^n)^{-2}(T_{[nt]}^n)^2\}_n$  is uniformly integrable for  $t \in (0, 1]$ , so by (3.1) and (2.10) we get

(3.11) 
$$\mathscr{L}((z_{[nt]}^n)^{-1}Z_{[nt]}^n) \xrightarrow{\mathsf{w}} \mathscr{N}(0, 1), \quad n \to +\infty.$$

From (3.11), (3.9) and the Theorem of Convergence of Types we get (3.8). Now observe that by assumption and (2.10) we have

$$\mathscr{L}(z_n^{-1}(S_n-\hat{S}_n)) \xrightarrow{\mathbf{w}} \mathscr{N}(0, 1), \quad n \to +\infty.$$

Thus, by Theorem 18.1.1 in [3] we have

(3.12) 
$$\frac{(\tau_{kn})^2}{(\tau_n)^2} \xrightarrow{n} k, \quad k \in \mathbb{N}, \ n \to +\infty.$$

Since

$$\begin{aligned} \mathscr{P}(|X_1 - \hat{X}_1| > \varepsilon z_n) &\leq \mathscr{P}(|X_1^n - \hat{X}_1^n| > 2^{-1}\varepsilon z_n) + 2\mathscr{P}(|X_1| \geq b_n) \\ &\leq 4\varepsilon^{-2}(z_1^n)^2 z_n^{-2} + n\mathscr{P}(|X_1| \geq b_n), \end{aligned}$$

so by (3.3) and (3.1) we obtain

$$\mathscr{L}(z_n^{-1}(S_{n+1}-\hat{S}_{n+1})) \xrightarrow{\mathbf{w}} \mathscr{N}(0, 1), \quad n \to +\infty.$$

Thus  $\lim_{n\to\infty} z_n z_{n+1}^{-1} = 1$ , so

(3.13) 
$$\tau_{n+1}\tau_n^{-1} \xrightarrow{n} 1, \quad n \to +\infty.$$

Let  $q \in \mathbb{N}$ ; then

$$\frac{(\tau_{q[nq^{-1}]})^2}{(\tau_{[nq^{-1}]})^2} \xrightarrow{n} q, \quad n \to +\infty.$$

But  $q[nq^{-1}] = n, n-1, ..., n-q-1$  and, by (3.13),

$$\frac{(\tau_n)^2}{(\tau_{\lfloor nq^{-1}\rfloor})^2} \xrightarrow{n} q, \quad n \to +\infty,$$

so by (3.12) we have

(3.15) 
$$\frac{(\tau_{[\omega n]})^2}{(\tau_n)^2} \xrightarrow{n} \omega, \quad n \to +\infty,$$

for every  $\omega$  rational. Let r be irrational and  $r \in (0, 1]$ ,  $c = r - \omega > 0$ . We show, following Peligrad [7], that

(3.16) 
$$\frac{(\tau_{[rn]})^2}{(\tau_{\cdot})^2} \xrightarrow{n} r, \quad n \to +\infty.$$

From (2.2) we have

$$|\tau_{[\omega n]}^n - \tau_{[rn]}^n| \leqslant \tau_{[rn]-[\omega n]}^n \leqslant (1 - \varrho_p)^{-1/2} (\tau_{[n(r-\omega)]+2}^n + 2\tau_1^n),$$

so taking lim sup over both sides we have, by (3.3),

$$\limsup_{n\to\infty} \tau_n^{-1} |\tau_{[mn]}^n - \tau_{[rn]}^n| \leqslant (1-\varrho_p)^{-1/2} \limsup_{n\to\infty} \tau_n^{-1} \tau_{[n(r-\omega)]}^n.$$

Now, it remains to show that the right-hand side disappears when  $\omega \nearrow r$ . We have

$$\frac{\tau_{[nc]}^n}{\tau_n} = \frac{\tau_{[n/2]}^n}{\tau_n} \times \frac{\tau_{[n/2]}^n}{\tau_{[n/2]}^n} \times \frac{\tau_{[n/2^3]}^n}{\tau_{[n/2^2]}^n} \times \frac{\tau_{[n/2^3]}^n}{\tau_{[n/2^3]}^n} \times \dots \times \frac{\tau_{[nc]}^n}{\tau_{[n/2^{1-\log c/\log 2]}}^n}.$$

Note that  $\limsup$  of the last multiplier is bounded by  $(1-\varrho_p)^{-1/2}$ , so

$$\limsup_{n \to \infty} \frac{\tau_{[n(r-\omega)]}^n}{\tau_n} \leq (1 - \varrho_p)^{-1/2} 2^{-(1/2)([-\log c/\log 2] - 1)} \leq K(r - \omega),$$

where K is a constant depending on  $\varrho_p$  only, i.e., (3.16) holds. By (3.8) and (3.16), for every  $r \in (0, 1]$  we have

$$\frac{(\tau_{[rn]})^2}{(\tau_n)^2} \xrightarrow{n} r, \quad n \to +\infty,$$

so by Theorem 1.3 in [10] the above holds for every r > 0, i.e.,  $\{(\tau_n)^2\}_n$  forms a regularly varying sequence with index 1.

The converse half. We have

$$\begin{split} \varphi_1 + \max_{1 \leq j \leq n} \mathscr{P}(|Z_n - Z_j^n| > z_n x_0) &\leq \varphi_1 + \max_{1 \leq j \leq n} \mathscr{P}(|Z_n - Z_j| > 2^{-1} z_n x_0) \\ &+ \max_{1 \leq j \leq N_A} \mathscr{P}(|Z_j - Z_j^n| > 2^{-1} z_n x_0) + \max_{N_O < j \leq n} \mathscr{P}(|Z_j - Z_j^n| > 2^{-1} z_n x_0), \end{split}$$

where  $N_{\delta}$  is such that  $\mathscr{P}(\tau_n^{-1}|R_n| > 2^{-1}x_0) \leqslant n\mathscr{P}(|X_1| > b_n) \leqslant \delta$  for  $n > N_{\delta}$ . The right-hand side of the above inequality can be estimated by

$$\varphi_1 + \frac{8}{x_0^2} \left( 1 + \max_{1 \le i \le n} \frac{(\tau_i)^2}{(\tau_n)^2} \right) + o(1) + \delta,$$

i.e., there exists  $N_0 = N(\delta, \varphi_1)$  such that for  $n \ge N_0$  and sufficiently large  $x_0$ 

$$\varphi_1 + \max_{1 \leq i \leq n} \mathcal{P}(|Z_n - Z_j^n| > z_n x_0) \leq \eta < 1,$$

since  $\max_{1 \le j \le n} \tau_j \tau_n^{-1}$  is bounded, by (3.4). Using (2.3), for  $n \ge N_0$ ,  $x \ge x_0$  we obtain

$$(3.17) \qquad \mathscr{P}(\max_{1 \leq i \leq n} |Z_i^n| > 2xz_n) \leqslant (1-\eta)^{-1} \mathscr{P}(|Z_n| > xz_n),$$

and since

$$\mathscr{P}(\max_{1 \leq i \leq n} |U_i^n| > x) \leq 2\mathscr{P}(\max_{1 \leq i \leq n} |Z_i^n| > 2^{-1}x),$$

so, by (3.17),  $\{\max_{1 \le i \le n} z_n^{-2} (U_i^n)^2\}_n$  is uniformly integrable. By the proof of Theorem 1 in [5] we have

(3.18) 
$$\max_{1 \le i \le n} \tau_n^{-1} |X_i^n| \xrightarrow{\mathscr{P}} 0, \quad n \to +\infty,$$

so for  $\mu_n = \text{med}(\max_{1 \le i \le n} \tau_n^{-1} |X_i^n|)$  we obtain

$$\mu_n \xrightarrow{n} 0, \quad n \to +\infty.$$

Thus

$$\mathscr{P}(\max_{1\leqslant i\leqslant n}z_n^{-1}|X_i^n|\geqslant x)\leqslant \mathscr{P}(|\max_{1\leqslant i\leqslant n}z_n^{-1}|X_i^n|-\mu_n|\geqslant x-\mu_n)$$

$$\leq 2\mathscr{P}(|\max_{1\leq i\leq n} z_n^{-1}|X_i^n| - \max_{1\leq i\leq n} z_n^{-1}|\hat{X}_i^n|| \geqslant x - \mu_n) \leq 4\mathscr{P}(\max_{1\leq i\leq n} z_n^{-1}|U_i^n| \geqslant x - \mu_n).$$

From this, (3.19), (3.18) and the uniform integrability of  $\{\max_{1 \le i \le n} z_n^{-2} (U_i^n)^2\}_n$  the equality (3.3) holds true.

Proof of Corollary 1. By (2.6), (3.5), (2.4) it suffices to prove that

$$\{b_n^{-2} \max_{1 \le i \le n} (X_i^* I(|X_i^*| < b_n))^2\}_n$$

is uniformly integrable, but this follows easily from the iid case.

Proof of Corollary 2. Under the assumptions of the corollary Peligrad [7] proved that for every  $k \in \mathbb{N}$ :

$$\frac{k^2 a_n^2}{\sigma^2(ka_n)} \xrightarrow{n} 0, \quad n \to +\infty,$$

where

$$\sigma^{2}(ka_{n}) = \operatorname{Var}\left(\sum_{i=1}^{n} X_{i}I(|X_{i}| < ka_{n}) - \mathscr{E}X_{i}I(|X_{i}| < ka_{n})\right),$$

and  $\{a_n\}_n$  is defined in (2.8). So there exists  $\{r_n\}_n$ ,  $\lim_n r_n = +\infty$ , such that, for

every  $\{x_n\}_n$ ,  $\lim_n x_n = +\infty$  and  $x_n = o(r_n)$ ,

(3.20) 
$$\frac{x_n^2 a_n^2}{\sigma^2(x_n a_n)} \xrightarrow{n} 0, \quad n \to +\infty.$$

On the other hand, by Theorem 1.1 in [10], there exists  $\{r'_n\}_n$ ,  $\lim_n r'_n = +\infty$ , such that, for every  $\{x_n\}_n$ ,  $\lim_n x_n = +\infty$  and  $x_n = o(r'_n)$ ,

$$(3.21) nx_n^2 \mathscr{P}(|X_1| > x_n a_n) \xrightarrow{n} 1, \quad n \to +\infty.$$

Now let  $b_n = x_n a_n$ , where  $\lim_n x_n = +\infty$ ,  $x_n = o(r_n \wedge r'_n)$ , and  $\tau_n = \sigma(x_n a_n)$ ; then (3.1)–(3.3) are fulfilled, so (3.4) holds. Observe that by (2.9) we have

$$\begin{split} \frac{[nt]}{\tau_n} \left| \mathscr{E} X_1 I(|X_1| > b_n) \right| &\leq \frac{[nt]}{\tau_n} \, \mathscr{E} |X_1| I(|X_1| > b_n) \\ &\sim 2 \, \frac{[nt]}{\sigma(x,a)} \, x_n a_n \mathscr{P}(|X_1| > x_n a_n), \quad n \to +\infty, \end{split}$$

so this and (3.20), (3.21) give (3.6). Since  $\tau_n \sim \sqrt{\pi/2} \,\mathscr{E} |T_n|$  and

$$\left|\frac{\mathscr{E}|S_n|-\mathscr{E}|T_n|}{\mathscr{E}|T_n|}\right| \leq \frac{n\mathscr{E}|X_1|I(|X_1|>b_n)}{\mathscr{E}|T_n|} \sim \frac{2nb_n\mathscr{P}(|X_1|>b_n)}{\sqrt{2/\pi}\tau_n}, \quad n \to +\infty,$$

so, as above, (3.7) holds.

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Remark. There are strictly stationary random sequences with infinite variance,  $\varphi$ -mixing, satisfying CLT and not satisfying WIP (i.e. (3.6)). As an example one can use a 1-dependent sequence in Example 2 of [6]. For this sequence, (3.5) does not hold.

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