

CONDITIONED RANDOM WALKS WITH RANDOM INDICES

BY

A. SZUBARGA AND D. SZYNAL (LUBLIN)

Abstract. Let $\{X_k, k \geq 1\}$ be a sequence of i.i.d. random variables with $EX_1 = 0$, $EX_1^2 = \sigma^2 < \infty$, and let $\{N_m, m \geq 0\}$, $N_0 = 0$ a.s., be a sequence of positive integer-valued random variables. Let $\{S_n, n \geq 0\}$ and $\{S_{N_m}, m \geq 0\}$ be defined by $S_0 = 0$ a.s., $S_n = X_1 + \dots + X_n$, $n \geq 1$, $S_{N_0} = 0$ a.s., $S_{N_m} = X_1 + X_2 + \dots + X_{N_m}$, $m \geq 1$. Put

$$N = \inf\{n: S_n < 0\}, \quad M = \max\{S_n: n \leq N\}.$$

In this note, under additional conditions on sequences $\{X_k, k \geq 1\}$ and $\{N_m, m \geq 0\}$, we investigate the limit behaviour of $P[M/\sigma\sqrt{N_m} \leq v | N > N_m]$, $P[\max_{0 \leq k \leq N_m} S_k/\sigma\sqrt{N_m} \leq v | N > N_m]$, and $P[N > N_m | M > v\sigma\sqrt{N_m}]$, where $v > 0$.

1. Introduction. Let $\{X_k, k \geq 1\}$ be a sequence of independent identically distributed random variables with $EX_1 = 0$, $0 < EX_1^2 = \sigma^2 < \infty$, and let $\{S_n, n \geq 0\}$ with $S_0 = 0$ and $S_n = X_1 + \dots + X_n$, $n \geq 1$, denote the random walk. Put $N = \inf\{n: S_n < 0\}$, $M = \max\{S_n: n \leq N\}$, and write $\mu = -ES_N$. The following result is known:

THEOREM 1. *Under the above assumptions:*

- (i)
$$\lim_{x \rightarrow \infty} xP[M > x] = \mu;$$
- (ii)
$$\lim_{x \rightarrow \infty} P[\sigma^2 N/2x^2 \leq u | M > x] = (2/\sqrt{\pi u}) \sum_{k=1}^{\infty} \exp(-k^2/u), \quad 0 < u < \infty;$$
- (iii)
$$\lim_{n \rightarrow \infty} P[M/\sqrt{n\sigma} \leq v | N > n] = 1 - v^{-1} \sqrt{\pi/2} + 2 \sum_{k=1}^{\infty} \exp(-2k^2 v^2);$$
- (iv)
$$\lim_{n \rightarrow \infty} P[N > n | M > \sigma\sqrt{nv}] = 2\sqrt{2/\pi} v \sum_{k=1}^{\infty} \exp(-2k^2 v^2), \quad 0 < v < \infty.$$

Proof. Statements (i)–(iii) have been proved in [2]. Putting now, in (ii), $n = [2ux^2/\sigma^2]$, $u = 1/(2v^2)$, we have

$$\begin{aligned} (1) \quad & \lim_{n \rightarrow \infty} (1 - P[N > n | M > \sigma \sqrt{nv}]) \\ &= \lim_{n \rightarrow \infty} (1 - P[S_1 \geq 0, \dots, S_n \geq 0 | M > \sigma \sqrt{nv}]) \\ &= 1 - 2\sqrt{2/\pi}v \sum_{k=1}^{\infty} \exp(-2k^2 v^2), \quad 0 < v < \infty, \end{aligned}$$

which proves (iv).

Moreover, we have

THEOREM 2. Under the above assumptions, for $x > 0$,

$$(2) \quad \lim_{n \rightarrow \infty} P[\max_{1 \leq k \leq n} S_k/\sigma \sqrt{n} \leq x | N > n] = \sum_{k=-\infty}^{+\infty} (-1)^k \exp(-k^2 x^2/2).$$

Proof. It is known that

$$(S_{[n \cdot y]}/\sigma \sqrt{n} | N > n) \Rightarrow W^+, \quad n \rightarrow \infty,$$

where W^+ is a Brownian meander (see [1] and [4]). In [3] (Corollary (2.2)) it has been proved that, for $x > 0$,

$$P[\sup_{0 \leq s \leq 1} W^+(s) \leq x] = \sum_{k=-\infty}^{+\infty} (-1)^k \exp(-k^2 x^2/2).$$

Hence, we conclude that, for $x > 0$,

$$\lim_{n \rightarrow \infty} P[\max_{1 \leq k \leq n} S_k/\sigma \sqrt{n} \leq x | N > n] = \sum_{k=-\infty}^{+\infty} (-1)^k \exp(-k^2 x^2/2),$$

which completes the proof of (2).

Let now $\{N_m, m \geq 1\}$ be a sequence of positive integer-valued random variables and put

$$S_{N_m} = X_1 + X_2 + \dots + X_{N_m}, \quad m \geq 1.$$

We are interested in the asymptotic behaviour of

$$P[M/\sigma \sqrt{N_m} \leq v | N > N_m], \quad 0 < v < \infty,$$

$$P[\max_{1 \leq k \leq N_m} S_k/\sigma \sqrt{N_m} \leq v | N > N_m],$$

and

$$P[S_1 \geq 0, \dots, S_{N_m} \geq 0 | M > v\sigma \sqrt{N_m}] = P[N > N_m | M > v\sigma \sqrt{N_m}].$$

2. Results. We now prove the following

THEOREM 3. Let $\{X_k, k \geq 1\}$ be a sequence of i.i.d. random variables with $EX_1 = 0, 0 < EX_1^2 = \sigma^2 < \infty$. Suppose that $\{N_m, m \geq 0\}, N_0 = 0$ a.s., is a sequence of positive integer-valued random variables independent of $\{X_k, k \geq 1\}$ and $\{\alpha_m, m \geq 1\}$ is a sequence of positive integers with $\alpha_m \rightarrow \infty, m \rightarrow \infty$, such that

$$(3) \quad \inf_m P[N_m \leq \alpha_m] = d \text{ for some positive } a, \text{ where } d > 0,$$

and

$$(4) \quad \sqrt{\alpha_m} P[N_m = n] \rightarrow 0, \quad m \rightarrow \infty, n \geq 1.$$

Then for $0 < v < \infty$

$$(5) \quad \lim_{m \rightarrow \infty} P[M/\sigma \sqrt{N_m} \leq v | S_1 \geq 0, \dots, S_{N_m} \geq 0] \\ = 1 - v^{-1} \sqrt{\pi/2} + 2 \sum_{k=1}^{\infty} \exp(-2k^2 v^2),$$

$$(6) \quad \lim_{m \rightarrow \infty} P[\max_{1 \leq k \leq N_m} S_k/\sigma \sqrt{N_m} \leq v | S_1 \geq 0, \dots, S_{N_m} \geq 0] \\ = \sum_{k=-\infty}^{+\infty} (-1)^k \exp(-k^2 v^2/2),$$

and

$$(7) \quad \lim_{m \rightarrow \infty} P[S_1 \geq 0, \dots, S_{N_m} \geq 0 | M > \sigma \sqrt{N_m} v] \\ = 2 \sqrt{2/\pi} v \sum_{k=1}^{\infty} \exp(-2k^2 v^2).$$

Proof. Assuming that, in (3), $a = 1$, we prove now (5) and (6). Note that

$$P[M/\sigma \sqrt{N_m} \leq v | S_1 \geq 0, \dots, S_{N_m} \geq 0] \\ = \sum_{n=1}^{\infty} t_{m,n} P[M/\sigma \sqrt{n} \leq v | S_1 \geq 0, \dots, S_n \geq 0],$$

where

$$t_{m,n} = P[S_1 \geq 0, \dots, S_n \geq 0] P[N_m = n] / P[S_1 \geq 0, \dots, S_{N_m} \geq 0], \\ n \geq 1, m \geq 1,$$

and

$$P[\max_{1 \leq k \leq N_m} S_k/\sigma \sqrt{N_m} \leq v | S_1 \geq 0, \dots, S_{N_m} \geq 0] \\ = \sum_{n=1}^{\infty} t_{m,n} P[\max_{1 \leq k \leq n} S_k/\sigma \sqrt{n} \leq v | S_1 \geq 0, \dots, S_n \geq 0].$$

Now taking into account that (Theorem 3.5, [7])

$$(8) \quad P[S_1 \geq 0, \dots, S_n \geq 0] \sim c/\sqrt{n}, \quad n \rightarrow \infty,$$

where

$$c = \exp \left\{ \sum_{k=1}^{\infty} (1/k)(1/2 - P[S_k > 0]) \right\},$$

we see, by assumption (3), that

$$(9) \quad \begin{aligned} P[S_1 \geq 0, \dots, S_{N_m} \geq 0] &\geq P[S_1 \geq 0, \dots, S_{N_m} \geq 0, N_m \leq \alpha_m] \\ &\geq P[S_1 \geq 0, \dots, S_{\alpha_m} \geq 0] P[N_m \leq \alpha_m] \geq (d \cdot c)/\sqrt{\alpha_m}, \quad m \rightarrow \infty. \end{aligned}$$

Hence, using assumption (4), we have

$$0 \leq t_{m,n} \leq \sqrt{\alpha_m} P[N_m = n]/(dc) \rightarrow 0, \quad m \rightarrow \infty.$$

Since

$$\sum_{n=1}^{\infty} t_{m,n} = 1, \quad m \geq 1,$$

$\{t_{m,n}\}$, $m \geq 1$, $n \geq 1$, is a Toeplitz matrix ([6], p. 472). Therefore, by (iii) and (2), we get (5) and (6), respectively.

Similarly arguing, we get

$$(10) \quad \begin{aligned} P[S_1 \geq 0, \dots, S_{N_m} \geq 0 | M > \sigma \sqrt{N_m} v] \\ = \sum_{n=1}^{\infty} c_{m,n} P[S_1 \geq 0, \dots, S_n \geq 0 | M > \sigma \sqrt{nv}], \end{aligned}$$

where

$$\begin{aligned} c_{m,n} &= P[M > \sigma \sqrt{nv}] P[N_m = n] / P[M > \sqrt{N_m} \sigma v], \quad m \geq 1, n \geq 1, \\ \sum_{n=1}^{\infty} c_{m,n} &= 1, \quad m \geq 1. \end{aligned}$$

Using assumption (3), we get

$$\begin{aligned} P[M > \sigma \sqrt{N_m} v] &\geq \sum_{\{n: N_m \leq \alpha_m\}} P[M > \sigma \sqrt{N_m} v, N_m = n] \\ &\geq P[M > \sigma \sqrt{\alpha_m} v] P[N_m \leq \alpha_m] \geq P[M > \sigma \sqrt{\alpha_m} v] \cdot d \end{aligned}$$

for sufficiently large m . Hence, by (4) and (i), we have

$$\begin{aligned} 0 \leq c_{m,n} &\leq P[N_m = n] / (P[M > \sigma \sqrt{\alpha_m} v] d) \\ &= \sigma v \sqrt{\alpha_m} P[N_m = n] / (d \sigma \sqrt{\alpha_m} P[M > \sqrt{\alpha_m} v \sigma]) \rightarrow 0, \quad m \rightarrow \infty. \end{aligned}$$

These facts together with (iv) imply (7).

COROLLARY 1. Suppose that a sequence $\{X_k, k \geq 1\}$ of i.i.d. random variables satisfies the assumptions of Theorem 3. If $\{N_m, m \geq 0\}$, $N_0 = 0$ a.s., is a sequence of positive integer-valued random variables independent of $\{X_k, k \geq 1\}$, and $\{\alpha_m, m \geq 1\}$ is a sequence of positive integers with $\alpha_m \rightarrow \infty, m \rightarrow \infty$, such that, for any given $\varepsilon > 0$,

$$(11) \quad P[|N_m/\alpha_m - \lambda| \geq \varepsilon] = o(1/\sqrt{\alpha_m}), \quad m \rightarrow \infty,$$

where λ is a random variable such that there exists an $a > 0$ such that

$$(12) \quad P[\lambda > a] = 1,$$

then (5), (6) and (7) hold true.

Proof. It is obvious that we can find $\{\alpha_m, m \geq 1\}$ such that (3) holds. For any fixed $n \in N$ assumptions (11) and (12) imply that for any given $\varepsilon, 0 < \varepsilon < a$, and sufficiently large m

$$\begin{aligned} \sqrt{\alpha_m} P[N_m = n] &\leq \sqrt{\alpha_m} P[N_m \leq \alpha_m(a - \varepsilon)] \leq \sqrt{\alpha_m} P[N_m \leq \alpha_m(\lambda - \varepsilon)] \times \\ &\quad \times \sqrt{\alpha_m} P[|N_m/\alpha_m - \lambda| \geq \varepsilon]. \end{aligned}$$

Hence, we have (4), which completes the proof of Corollary 1.

The following example shows that assumption (11), in general, cannot be weakened in that sense that o can not be replaced by O , as it is in the random central limit theorem.

Example 1. Let $\{N_m, m \geq 5\}$ be a sequence of positive integer valued random variables such that $P[N_m = 1] = 1/\sqrt{m}$, $P[N_m = 2] = 1/\sqrt{m}$, $P[N_m = m] = 1 - 2/\sqrt{m}$, $m \geq 5$.

Suppose that $\{X_k, k \geq 1\}$ is a sequence of Theorem 3 independent of $\{N_m, m \geq 5\}$. We see that $N_m/m \xrightarrow{P} 1, m \rightarrow \infty$, and, for $\varepsilon, 0 < \varepsilon < 1/4$, $P[|N_m/m - 1| \geq \varepsilon] = 2/\sqrt{m}, m \geq 5$.

Moreover, by (2) and (8), one can verify that, for $v > 0$,

$$\begin{aligned} &P\left[\max_{1 \leq k \leq N_m} S_k/\sigma \sqrt{N_m} \leq v \mid S_1 \geq 0, \dots, S_{N_m} \geq 0\right] \\ &\rightarrow (P[S_1 \leq \sigma v, S_1 \geq 0] + P[\max(S_1, S_2) \leq \sigma \sqrt{2}v, S_1 \geq 0, S_2 \geq 0]) + \\ &+ c \sum_{k=-\infty}^{+\infty} (-1)^k \exp(-k^2 v^2/2) / (P[S_1 \geq 0] + P[S_1 \geq 0, S_2 \geq 0] + c) \\ &\neq \sum_{k=-\infty}^{+\infty} (-1)^k \exp(-k^2 v^2/2). \end{aligned}$$

Similarly, one can verify that (5) and (7) do not hold with the considered sequence $\{N_m, m \geq 5\}$.

Consider now the case where $\{X_k, k \geq 1\}$ and $\{N_m, m \geq 0\}$ are not independent. We are able to prove the following

THEOREM 4. *Let $\{X_k, k \geq 1\}$ be a sequence of i.i.d. random variables with $EX_1 = 0$, $0 < EX_1^2 = \sigma^2 < \infty$. Suppose that $\{N_m, m \geq 0\}$, $N_0 = 0$ a.s., is a sequence of positive integer-valued random variables, and $\{\alpha_m, m \geq 1\}$ is a sequence of positive numbers with $\alpha_m \rightarrow \infty$, $m \rightarrow \infty$, such that, for any given $\varepsilon > 0$,*

$$(13) \quad P[a - \varepsilon \leq N_m/\alpha_m \leq b + \varepsilon] = o(1/\sqrt{\alpha_m}),$$

where $0 < a \leq b < \infty$ are constant.

Then, for $v > 0$,

$$(14) \quad \begin{aligned} & \sqrt{a/b} \{1 - (va/b)^{-1} \sqrt{\pi/2} + 2 \sum_{k=1}^{\infty} \exp(-2k^2(va/b)^2)\} \\ & \leq \liminf_{m \rightarrow \infty} P[M \leq v\sigma \sqrt{N_m} | S_1 \geq 0, \dots, S_{N_m} \geq 0] \\ & \leq \limsup_{m \rightarrow \infty} P[M \leq v\sigma \sqrt{N_m} | S_1 \geq 0, \dots, S_{N_m} \geq 0] \\ & \leq \sqrt{b/a} \{1 - (vb/a)^{-1} \sqrt{\pi/2} + 2 \sum_{k=1}^{\infty} \exp(-2k^2(vb/a)^2)\}, \end{aligned}$$

$$(15) \quad \begin{aligned} & \sqrt{a/b} \sum_{k=-\infty}^{+\infty} \{\exp(-k^2(va/b)^2/2)\} (-1)^k \\ & \leq \liminf_{m \rightarrow \infty} P[\max_{1 \leq k \leq N_m} S_k/\sigma \sqrt{N_m} \leq v | S_1 \geq 0, \dots, S_{N_m} \geq 0] \\ & \leq \limsup_{m \rightarrow \infty} P[\max_{1 \leq k \leq N_m} S_k/\sigma \sqrt{N_m} \leq v | S_1 \geq 0, \dots, S_{N_m} \geq 0] \\ & \leq \sqrt{b/a} \sum_{k=-\infty}^{+\infty} \{\exp(-k^2(vb/a)^2/2)\} (-1)^k, \end{aligned}$$

and

$$(16) \quad \begin{aligned} & \sqrt{a/b} 2 \sqrt{2/\pi} v \sum_{k=1}^{\infty} \exp(-2k^2 v^2) \\ & \leq \liminf_{m \rightarrow \infty} P[S_1 \geq 0, \dots, S_{N_m} \geq 0 | M > v\sigma \sqrt{N_m}] \\ & \leq \limsup_{m \rightarrow \infty} P[S_1 \geq 0, \dots, S_{N_m} \geq 0 | M > v\sigma \sqrt{N_m}] \\ & \leq \sqrt{b/a} 2 \sqrt{2/\pi} v \sum_{k=1}^{\infty} \exp(-2k^2 v^2). \end{aligned}$$

Proof. Put $A_m = \{n, (a-\varepsilon)\alpha_m \leq n \leq (b+\varepsilon)\alpha_m\}$. Then, by (13) and (8), we have

$$\begin{aligned} r_m &= P[S_1 \geq 0, \dots, S_{N_m} \geq 0] \\ &\leq P[S_1 \geq 0, \dots, S_{N_m} \geq 0, N_m \in A_m] + P[N_m \in A_m^c] \\ &\leq c/\sqrt{[(a-\varepsilon)\alpha_m]} + o(1/\sqrt{\alpha_m}). \end{aligned}$$

Similarly, we get

$$\begin{aligned} r_m &\geq P[S_1 \geq 0, \dots, S_{[(b+\varepsilon)\alpha_m]} \geq 0] - P[N_m \in A_m^c] \\ &\geq c/\sqrt{[(b+\varepsilon)\alpha_m]} - o(1/\sqrt{\alpha_m}). \end{aligned}$$

Hence, we obtain

$$(17) \quad -o(1/\sqrt{\alpha_m}) + c/\sqrt{[(b+\varepsilon)\alpha_m]} \leq r_m \leq c/\sqrt{[(a-\varepsilon)\alpha_m]} + o(1/\sqrt{\alpha_m}).$$

Thus

$$(18) \quad \lim_{m \rightarrow \infty} P[N_m \in A_m^c]/r_m = 0.$$

Hence, to prove (14), it is enough to consider

$$P[M \leq v\sigma\sqrt{N_m}, S_1 \geq 0, \dots, S_{N_m} \geq 0, N_m \in A_m]/r_m.$$

Note that

$$\begin{aligned} &P[M \leq v\sigma\sqrt{[(b+\varepsilon)\alpha_m]}, S_1 \geq 0, \dots, S_{[(a-\varepsilon)\alpha_m]} \geq 0, N_m \in A_m] \\ &\leq P[M \leq v\sigma\sqrt{[(b+\varepsilon)\alpha_m]/[(a-\varepsilon)\alpha_m]} \sqrt{[(a-\varepsilon)\alpha_m]}, S_1 \geq 0, \dots, S_{[(a-\varepsilon)\alpha_m]} \geq 0]. \end{aligned}$$

Moreover, we see that for any given $\eta > 0$ and sufficiently large m ,

$$\sqrt{[(b+\varepsilon)\alpha_m]/[(a-\varepsilon)\alpha_m]} \leq \sqrt{(b+\varepsilon)/(a-\varepsilon)} + \eta.$$

Hence, we get

$$\begin{aligned} &P[M \leq v\sigma\sqrt{N_m}, S_1 \geq 0, \dots, S_{N_m} \geq 0]/r_m \\ &\leq (P[S_1 \geq 0, \dots, S_{[(a-\varepsilon)\alpha_m]} \geq 0]/r_m) \times \\ &\quad \times P[M \leq v\sigma(\sqrt{(b+\varepsilon)/(a-\varepsilon)} + \eta)\sqrt{[(a-\varepsilon)\alpha_m]} | S_1 \geq 0, \dots, S_{[(a-\varepsilon)\alpha_m]} \geq 0]. \end{aligned}$$

Note that, for any given $\delta > 0$ and sufficiently large m ,

$$P[S_1 \geq 0, \dots, S_{[(a-\varepsilon)\alpha_m]} \geq 0]/r_m \leq \sqrt{(b+\varepsilon)/(a-\varepsilon)} + \delta.$$

Therefore, for sufficiently large m ,

$$(19) \quad \begin{aligned} &P[M \leq v\sigma \sqrt{N_m}, S_1 \geq 0, \dots, S_{N_m} \geq 0, N_m \in A_m]/r_m \\ &\leq P[M \leq v\sigma (\sqrt{(b+\varepsilon)/(c-\varepsilon)} + \eta) \sqrt{[(a-\varepsilon)\alpha_m]} | S_1 \geq 0, \\ &\quad \dots, S_{[(a-\varepsilon)\alpha_m]} \geq 0] (\sqrt{(b+\varepsilon)/(a-\varepsilon)} + \delta). \end{aligned}$$

Similarly we get, for any given $\Delta > 0$ and sufficiently large m ,

$$(20) \quad \begin{aligned} &P[M \leq v\sigma (\sqrt{(a-\varepsilon)/(b+\varepsilon)} - \Delta) \sqrt{[(b+\varepsilon)\alpha_m]} | S_1 \geq 0, \dots, S_{[(b+\varepsilon)\alpha_m]} \geq 0] \times \\ &\quad \times (\sqrt{(a-\varepsilon)/(b+\varepsilon)} - \Delta) - o(1/\sqrt{\alpha_m}) \\ &\leq P[M \leq v\sigma \sqrt{N_m}, S_1 \geq 0, \dots, S_{N_m} \geq 0, N_m \in A_m]/r_m. \end{aligned}$$

Letting now $m \rightarrow \infty$, $\eta \rightarrow 0$, $\delta \rightarrow 0$, $\Delta \rightarrow 0$, $\varepsilon \rightarrow 0$, we get (14). In the similar way one can get (15).

We now prove (16). Note that, for any given $\varepsilon > 0$,

$$P[M > v\sigma \sqrt{N_m}] \geq P[M > v\sigma \sqrt{[(b+\varepsilon)\alpha_m]}] - P[N_m \in A_m^c],$$

and

$$P[M > v\sigma \sqrt{N_m}] \leq P[M > v\sigma \sqrt{[(a-\varepsilon)\alpha_m]}] + P[N_m \in A_m^c].$$

Moreover, we see that

$$\begin{aligned} &P[S_1 \geq 0, \dots, S_{N_m} \geq 0 | M > v\sigma \sqrt{N_m}] \\ &= P[S_1 \geq 0, \dots, S_{N_m} \geq 0, M > v\sigma \sqrt{N_m}] / P[M > v\sigma \sqrt{N_m}] \\ &\leq \frac{P[S_1 \geq 0, \dots, S_{[(a-\varepsilon)\alpha_m]} \geq 0, M > \sigma v \sqrt{[(a-\varepsilon)\alpha_m]}] + P[N_m \in A_m^c]}{P[M > v\sigma \sqrt{N_m}]} \\ &\leq \frac{P[N_m \in A_m^c] + P[S_1 \geq 0, \dots, S_{[(a-\varepsilon)\alpha_m]} \geq 0 | M > v\sigma \sqrt{[(a-\varepsilon)\alpha_m]}]}{P[M > v\sigma \sqrt{[(b+\varepsilon)\alpha_m]}] - P[N_m \in A_m^c]} \times \\ &\quad \times P[M > v\sigma \sqrt{[(a-\varepsilon)\alpha_m]}] \end{aligned}$$

and

$$\begin{aligned} &P[S_1 \geq 0, \dots, S_{N_m} \geq 0 | M > v\sigma \sqrt{N_m}] \\ &\geq \frac{-P[N_m \in A_m^c] + P[M > \sigma v \sqrt{[(b+\varepsilon)\alpha_m]}]}{P[M > v\sigma \sqrt{[(a-\varepsilon)\alpha_m]}] + P[N_m \in A_m^c]} \times \\ &\quad \times P[S_1 \geq 0, \dots, S_{[(b+\varepsilon)\alpha_m]} \geq 0 | M > v\sigma \sqrt{[(b+\varepsilon)\alpha_m]}]. \end{aligned}$$

Hence, by the assumptions and (i), we obtain (16).

COROLLARY 2. If (13) holds with $a = b$, then (5), (6), and (7) hold true.

Remark 1. Assumption (13) holds true if we assume, for example, that, for any given $\varepsilon > 0$,

$$(13) \quad P[|N_m/\alpha_m - \lambda| \geq \varepsilon] = o(1/\sqrt{\alpha_m}), \quad P[a \leq \lambda \leq b] = 1,$$

where λ is a random variable, and $0 < a \leq b < \infty$ are constants.

Acknowledgement. The authors wish to thank the referee for valuable remarks which led to a considerable simplification of the paper.

REFERENCES

- [1] E. Bolthausen, *On a functional central limit theorem for random walks conditioned to stay positive*, Ann. Probability 4 (1976), p. 480–485.
- [2] R. Doney, *Note on conditioned random walk*, J. Appl. Probability 20 (1983), p. 409–412.
- [3] R. Durrett and D. Iglehart, *Functions of Brownian meander and Brownian excursion*, Ann. Probability 5 (1977), p. 130–135.
- [4] D. Iglehart, *Functional central limit theorems for random walks conditioned to stay positive*, ibidem 2 (1974), p. 608–619.
- [5] D. Kennedy, *The distribution of the maximum of Brownian excursion*, J. Appl. Probability 13 (1976), p. 371–376.
- [6] A. Rényi, *Probability theory*, Amsterdam-London 1970.
- [7] F. Spitzer, *A Tauberian theorem and its probability interpretation*, Trans. Amer. Math. Soc. 94 (1960), p. 150–169.

Mathematical Institute
 Maria Curie-Skłodowska University
 ul. Nowotki 10
 20-031 Lublin, Poland

Received on 15. 8. 1985;
 revised version on 7. 1. 1986
