

ON THE NUMBER OF k -TREES IN A RANDOM GRAPH

BY

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Abstract. Let $K_{n,p}$ denote a random graph obtained from a complete labelled graph K_n on n vertices by independent deletion of its edges with the prescribed probability $q = 1 - p$, $0 < p < 1$. Moreover, let $p = p(n)$ and let $X_{n,r}^{(k)}$ denote the number of r -vertex subgraphs ($r \geq k + 1$) of a random graph $K_{n,p}$ being k -trees. In this paper we prove that, under some conditions imposed on probability $p(n)$ as $n \rightarrow \infty$, the random variable $X_{n,r}^{(k)}$ has asymptotically the Poisson or normal distribution. We generalize earlier results of Erdős and Rényi [2] dealing with the distribution of the number of trees (i.e. random variable $X_{n,r}^{(1)}$) as well as the results of Schürger [7] on the number of cliques in $K_{n,p}$ (i.e. random variable $X_{n,k+1}^{(k)}$).

1. Introduction. Let us consider a random graph (r.g.) $K_{n,p}$ obtained from a labelled complete graph K_n by means of the following procedure:

Each of $\binom{n}{2}$ edges of K_n is independently deleted with the prescribed probability q , $0 < q < 1$, i.e. each edge remains in K_n with probability $p = 1 - q$ and the expected number of edges of an r.g. $K_{n,p}$ equals $\binom{n}{2}p$.

Here we shall consider the asymptotic distribution of the number of subgraphs being k -trees of size r in an r.g. $K_{n,p(n)}$, where the edge probability p depends on the size n of an r.g. and $n \rightarrow \infty$.

The notion of k -tree ($k = 1, 2, \dots$) was introduced in [3] and can be defined either as a k -dimensional simplicial complex with certain properties or as a graph. We shall use here an inductive definition of k -tree (see [5] or [6]). A k -tree of size $k + 1$ is a complete graph on $k + 1$ vertices. A k -tree $T_{r+1}^{(k)}$ of size $r + 1$ ($r \geq k + 1$) is obtained from an arbitrary k -tree $T_r^{(k)}$ of size r by adding a new vertex and joining it to those k points of $T_r^{(k)}$

which form a complete graph. A k -tree of size r consists of $r-k$ complete subgraphs of size $k+1$, $k(r-k)+1$ complete subgraphs of size k and has $kr - \binom{k+1}{2}$ edges. It has been shown (see [1] and [4]) that the total number of k -trees which can be formed on r labelled vertices equals

$$\binom{r}{k} [k(r-k)+1]^{r-k-2}.$$

In this paper we generalize the results of Erdős and Rényi [2] on the number of trees and the results of Schürger [7] on cliques. The methods of proofs are mainly those of [7].

2. The Lemma. Suppose that $G_i = (V(G_i), E(G_i))$, $i = 1, 2, \dots, t$, are graphs whose vertex sets $V(G_i)$ as well as edge families $E(G_i) \subset V(G_i) \times V(G_i)$ are not necessarily disjoint. Then by their union $\bigcup_{i=1}^t G_i$ we mean a graph $G = (V(G), E(G))$, where

$$V(G) = \bigcup_{i=1}^t V(G_i) \quad \text{and} \quad E(G) = \bigcup_{i=1}^t E(G_i).$$

Denote by $|X|$ the cardinality of a set X and by a ,

$$a = a(r, k) = kr - \binom{k+1}{2},$$

the number of edges of a k -tree of size r . In this section we state the following purely graph-theoretic result:

LEMMA. Suppose that $T_{r,1}^{(k)}, T_{r,2}^{(k)}, \dots, T_{r,t}^{(k)}$ are pairwise different k -trees of size r , $r \geq k+1$, not all of which are pairwise vertex disjoint, $G_t = \bigcup_{i=1}^t T_{r,i}^{(k)}$ is their union, $|V(G_t)| = v_t$, and $|E(G_t)| = e_t$. Then $v_t/e_t \leq \max\{b_0, b_1, b_2\}$, where

$$(1) \quad b_0 = b_0(t) = \frac{(t-1)r}{(t-1)a+1}, \quad b_1 = b_1(t) = \frac{tr-1}{ta},$$

$$b_2 = b_2(t) = \frac{(t-1)r+1}{(t-1)a+k}.$$

Proof (by induction on t). Put

$$h_t = |V(T_{r,t}^{(k)} - \bigcup_{i=1}^{t-1} T_{r,i}^{(k)})|, \quad t \geq 2.$$

It is easy to check that if $1 \leq h_t \leq r-k$, then $e_t \geq e_{t-1} + kh_t$, whereas

$$e_t \geq e_{t-1} + a - \binom{r-h_t}{2} \quad \text{if } r-k+1 \leq h_t \leq r-1.$$

One can also check that the inequality

$$(2) \quad \frac{v_{t-1} + r - i}{e_{t-1} + a - \binom{i}{2}} \leq \frac{v_{t-1} + r - 1}{e_{t-1} + a}$$

holds for $i = 1, 2, \dots, k-1$. Moreover, it should be noticed that $b_0(t)$, $b_1(t)$, and $b_2(t)$ are increasing functions of t . Having these facts in mind we begin our proof by considering $t = 2$.

Case 1. Let $h_2 = 0$ ($r \geq k+2$). Then $v_2 = v_1 = r$, $e_2 \geq e_1 + 1$ and, consequently,

$$\frac{v_2}{e_2} \leq \frac{r}{a+1} = b_0(2).$$

Case 2. If $1 \leq h_2 \leq r-k$, then

$$\frac{v_2}{e_2} = \frac{r+h_2}{e_2} \leq \frac{r+h_2}{a+kh_2} \leq \frac{r+1}{a+k} = b_2(2).$$

Case 3. Let $r-k+1 \leq h_2 \leq r-1$. Now, by (2) we have

$$\frac{v_2}{e_2} \leq \frac{r+h_2}{e_1+a-\binom{r-h_2}{2}} \leq \frac{2r-1}{2a} = b_1(2)$$

and we arrive at the thesis for $t = 2$.

Let us assume that our thesis is true for some $t \geq 3$. Suppose that G_{t-1} is the union of $t-1$ k -trees of size r not all of which are pairwise vertex disjoint and

$$\frac{v_{t-1}}{e_{t-1}} \leq \max \{b_0(t-1), b_1(t-1), b_2(t-1)\}$$

holds. Let us take now the union of G_{t-1} with a k -tree also of size r .

Case 1. If $h_t = 0$, then $v_t/e_t \leq v_{t-1}/e_{t-1}$ and the thesis follows immediately.

Case 2. Let $1 \leq h_t \leq r-k$. Assume first that, for G_{t-1} , $b_0(t-1)$ exceeds $b_1(t-1)$ and $b_2(t-1)$. Then

$$\begin{aligned} \frac{v_t}{e_t} &\leq \frac{v_{t-1} + h_t}{e_{t-1} + kh_t} \leq \frac{(t-2)r\{v_{t-1} + h_t\}}{v_{t-1}\{(t-2)a+1\} + kh_2(t-2)r} \\ &= \frac{(t-2)r\{v_{t-1} + h_t\}}{\{(t-2)a+1\} \left\{ v_{t-1} + h_t \frac{kr(t-2)}{a(t-2)+1} \right\}} \leq \frac{(t-2)r}{(t-2)a+1} = b_0(t-1) \leq b_0(t). \end{aligned}$$

Similarly, if $b_1(t-1)$ is maximal, then

$$\frac{v_t}{e_t} \leq \frac{\{(t-1)r-1\} \{v_{t-1}+h_t\}}{a(t-1) \left\{ v_{t-1}+h_t \frac{k[(t-1)r-1]}{a(t-1)} \right\}} \leq b_1(t-1) \leq b_1(t),$$

and if, for G_{t-1} , $b_2(t-1)$ is greater than $b_0(t-1)$ and $b_1(t-1)$, then

$$\frac{v_t}{e_t} \leq b_2(t).$$

Case 3. Let $r-k+1 \leq h_t \leq r-1$. As before we assume first that, for G_{t-1} , $b_0(t-1)$ exceeds $b_1(t-1)$ and $b_2(t-1)$. Then by formula (2) and the induction assumption we obtain

$$\begin{aligned} \frac{v_t}{e_t} &\leq \frac{v_{t-1}+h_t}{e_{t-1}+a-\binom{r-h_t}{2}} \leq \frac{v_{t-1}+r-1}{e_{t-1}+a} \\ &\leq \frac{r(t-2)\{v_{t-1}+r-1\}}{\{a(t-2)+1\} \left\{ v_{t-1} + \frac{ar(t-2)}{a(t-2)+1} \right\}} \leq b_0(t-1) \end{aligned}$$

if $r \leq a(t-2)+1$, which is true for all $t \geq 3$ and $k \geq 1$. Therefore $v_t/e_t \leq b_0(t)$.

If $b_1(t-1)$ is maximal, then

$$\frac{v_t}{e_t} \leq \frac{\{r(t-1)+1\} \{v_{t-1}+r-1\}}{a(t-1) \{v_{t-1}+r-1/(t-1)\}} \leq b_1(t-1) \leq b_1(t).$$

Finally, if, for G_{t-1} , $b_2(t-1)$ is greater than $b_0(t-1)$ and $b_1(t-1)$, then

$$\frac{v_t}{e_t} \leq \frac{\{r(t-2)+1\} \{v_{t-1}+r-1\}}{\{a(t-2)+k\} \left\{ v_{t-1} + \frac{a[r(t-2)+1]}{a(t-2)+k} \right\}} \leq b_2(t-1)$$

whenever $a(t-1) \geq k(r-1)$, which holds for all $t \geq 3$. Consequently, $v_t/e_t \leq b_2(t)$.

To complete the proof of our lemma we consider also the situation when G_{t-1} consists of $t-1$ pairwise vertex disjoint k -trees of size r and next we form the union of G_{t-1} with some k -tree of size r .

Case 1. If $h_t = 0$ ($r \geq k+2$), then

$$\frac{v_t}{e_t} \leq \frac{v_{t-1}}{e_{t-1}+1} = \frac{r(t-1)}{a(t-1)+1} = b_0(t).$$

Case 2. If $1 \leq h_t \leq r-k$, then

$$\frac{v_t}{e_t} \leq \frac{r(t-1)+h_t}{a(t-1)+kh_t} \leq b_2(t).$$

Case 3. If $r-k+1 \leq h_t \leq r-1$, then

$$\frac{v_t}{e_t} \leq \frac{r(t-1)+h_t}{a(t-1)+a-\binom{r-h_t}{2}} \leq b_1(t).$$

Thus the proof is complete.

3. Asymptotic distribution of the number of *k*-trees. Let us denote by $X_{n,r}^{(k)}$ the number of *k*-trees of size *r* in an r.g. $K_{n,p(n)}$. We shall prove the following theorem:

THEOREM 1. Suppose that $r \geq k+1$, $k = 1, 2, \dots$, and that

$$(3) \quad \lim_{n \rightarrow \infty} p(n)n^{r/a} = \varrho \in (0, \infty) \text{ exists.}$$

Then

$$\lim_{n \rightarrow \infty} P(X_{n,r}^{(k)} = i) = \frac{\lambda^i}{i!} e^{-\lambda}, \quad i = 0, 1, 2, \dots,$$

where

$$(4) \quad \lambda = \frac{1}{r!} \varrho^a \binom{r}{k} [k(r-k)+1]^{r-k-2},$$

and

$$a = a(r, k) = kr - \binom{k+1}{2}.$$

Proof. Let $\mathcal{T}_r^{(k)}$ denote the family of all *k*-trees of size *r* which can be formed on the set of *n* labelled vertices $\{1, 2, \dots, n\}$. Suppose that $T_{r,0}^{(k)} \in \mathcal{T}_r^{(k)}$ and

$$I_1(T_{r,0}^{(k)}) = \begin{cases} 1 & \text{if } T_{r,0}^{(k)} \in K_{n,p(n)}, \\ 0 & \text{otherwise,} \end{cases}$$

whereas

$$I_2(T_{r,0}^{(k)}) = \begin{cases} 1 & \text{if } T_{r,0}^{(k)} \in K_{n,p(n)} \text{ and } T_{r,0}^{(k)} \text{ is vertex disjoint with} \\ & \text{all other } k\text{-trees of size } r \text{ contained in } K_{n,p(n)}, \\ 0 & \text{otherwise.} \end{cases}$$

Moreover, let

$$S_i^{(i,j)} = \sum^{(i)} E(I_j(T_{r,1}^{(k)}) I_j(T_{r,2}^{(k)}) \dots I_j(T_{r,i}^{(k)})), \quad i, j = 1, 2,$$

where $E(\cdot)$ denotes the expected value, and the summation is over all combinations of t ($t \geq 1$) different k -trees from the family $\mathcal{T}_r^{(k)}$ which for $i = 1$ are pairwise vertex disjoint and for $i = 2$ are not pairwise vertex disjoint. First, we show that if (3) holds, then $S_t^{(2,1)} \rightarrow 0$ as $n \rightarrow \infty$.

Assume that $t \geq 2$, $n \geq tr$, $T_{r,1}^{(k)}, T_{r,2}^{(k)}, \dots, T_{r,t}^{(k)}$ are different k -trees from $\mathcal{T}_r^{(k)}$ which are not pairwise vertex disjoint, and put

$$w_j = \left| \bigcup_{i=1}^{j-1} (V(T_{r,i}^{(k)}) \cap V(T_{r,j}^{(k)})) \right|.$$

Therefore

$$v_j = \left| V\left(\bigcup_{i=1}^j T_{r,i}^{(k)}\right) \right| = jr - w_2 - \dots - w_j, \quad j = 2, 3, \dots, t.$$

Thus, for every $r \geq k+1$ and given t we get

$$\begin{aligned} S_t^{(2,1)} &\leq \frac{1}{t!} \sum_{\substack{0 \leq w_2, \dots, w_t \leq r \\ 1 \leq w_2 + \dots + w_t \leq (t-1)r}} \binom{n}{r} \binom{r}{w_2} \binom{n-r}{r-w_2} \times \dots \times \\ &\times \binom{(t-1)r - w_2 - \dots - w_{t-1}}{w_t} \binom{n - (t-1)r + w_2 + \dots + w_{t-1}}{r - w_t} \times \\ &\times \left\{ \binom{r}{k} [k(r-k) + 1]^{r-k-2} \right\}^t p^{e_t}, \end{aligned}$$

where e_t is the number of edges of the graph $\bigcup_{i=1}^t T_{r,i}^{(k)}$. Consequently, we obtain

$$\begin{aligned} S_t^{(2,1)} &\leq \frac{1}{t! r!} \left\{ \binom{r}{k} [k(r-k) + 1]^{r-k-2} \right\}^t \times \\ &\times \sum n^{tr - w_2 - \dots - w_t} p^{e_t} \left\{ \prod_{j=1}^{t-1} \binom{jr - w_2 - \dots - w_j}{w_{j+1}} [(r - w_{j+1})!]^{-1} \right\}. \end{aligned}$$

By the Lemma we have

$$(5) \quad e_t \geq \frac{tr - w_2 - \dots - w_t}{\max\{b_0, b_1, b_2\}},$$

where b_0 , b_1 , and b_2 are determined by formula (1). It follows from (5) and (3) that for every $t \geq 2$

$$(6) \quad S_t^{(2,1)} = o(1).$$

On the other hand, for $t \geq 1$ we get

$$0 \leq S_t^{(1,1)} - S_t^{(1,2)} \leq (t+1) S_{t+1}^{(2,1)}.$$

Consequently, for $t \geq 2$ we have

$$S_t^{(1,2)} = S_t^{(1,1)} + o(1).$$

But

$$\begin{aligned} S_t^{(1,1)} &= \frac{1}{t!} \binom{n}{r} \binom{n-r}{r} \dots \binom{n-(t-1)r}{r} \left\{ \binom{r}{k} [k(r-k)+1]^{r-k-2} p^a \right\}^t \\ &\sim \frac{1}{t!} \left\{ \frac{1}{r!} (pn^{r/a})^a \binom{r}{k} [k(r-k)+1]^{r-k-2} \right\}^t. \end{aligned}$$

Therefore

$$(7) \quad \lim_{n \rightarrow \infty} S_t^{(1,2)} = \frac{\lambda^t}{t!}, \quad t \geq 1,$$

where λ is given by (4).

If we denote by $Y_{n,r}^{(k)}$ the number of all k -trees of size r being vertex disjoint with all other k -trees of size r in $K_{n,p(n)}$, then using (7) and Bonferroni's inequalities

$$\sum_{j=i}^{i+2s-1} (-1)^{j-i} \binom{j}{i} S_j^{(1,2)} \leq P(Y_{n,r}^{(k)} = i) \leq \sum_{j=i}^{i+2s} (-1)^{j-i} \binom{j}{i} S_j^{(1,2)}, \quad s \geq 1,$$

we obtain

$$\lim_{n \rightarrow \infty} P(Y_{n,r}^{(k)} = i) = \frac{\lambda^i}{i!} e^{-\lambda},$$

i.e. $Y_{n,r}^{(k)}$ has the Poisson distribution with the expectation λ . The thesis of our theorem follows immediately from the fact that, by (6),

$$P(X_{n,r}^{(k)} \neq Y_{n,r}^{(k)}) \rightarrow 0 \quad \text{as } n \rightarrow \infty.$$

Consider now two specific k -trees, namely a 1-tree and a k -tree of size $k+1$.

In the first case ($k=1$) the 1-tree is simply a tree in the usual sense and $X_{n,r}^{(1)}$ denotes the number of trees of size r in $K_{n,p(n)}$. Thus, from Theorem 1 we obtain the following result which was proved earlier by Erdős and Rényi [2]:

COROLLARY 1. Suppose that $r \geq 2$ and

$$(8) \quad \lim_{n \rightarrow \infty} p(n)n^{r/(r-1)} = \varrho \in (0, \infty) \text{ exists.}$$

Then the r.v. $X_{n,r}^{(1)}$ has asymptotically the Poisson distribution with the expectation

$$\lambda = \frac{1}{r!} \varrho^{r-1} r^{r-2}.$$

In fact, if condition (8) holds, then in an r.g. $K_{n,p(n)}$ all subgraphs of size r being trees are almost surely isolated (see [2], p. 27). Therefore, the r.v. $X_{n,r}^{(1)}$ is the number of isolated trees of size r .

In the second case, where $r = k+1$, the k -tree is the smallest one and is simply a complete graph on $k+1$ vertices. The random variable $X_{n,k+1}^{(k)}$ is now the number of complete subgraphs of size $k+1$ ($k+1$ - cliques) in an r.g. $K_{n,p(n)}$.

Thus, also from Theorem 1, we obtain the following result which was proved earlier by Schürger in [7]:

COROLLARY 2. *Suppose that $k \geq 1$ and*

$$\lim_{n \rightarrow \infty} p(n)n^{2/k} = \varrho \in (0, \infty) \text{ exists.}$$

Then the random variable $X_{n,k+1}^{(k)}$ has asymptotically the Poisson distribution with the expectation

$$\lambda = \frac{1}{r!} \varrho^{k(k+1)/2}.$$

Finally, basing on Theorem 1 and using the fact that if a random variable X_λ has the Poisson distribution with the expectation λ , then $(X - \lambda)/\lambda^{1/2}$ has the standardized normal distribution as $\lambda \rightarrow \infty$, one can deduce the following

THEOREM 2. *Let $r \geq k+1$ be fixed and suppose that*

$$\lim_{n \rightarrow \infty} p(n)n^{r/a} = \infty,$$

whereas

$$\lim_{n \rightarrow \infty} p(n)n^{r/a-\delta} = 0 \quad \text{for all } \delta > 0.$$

Then

$$\lim_{n \rightarrow \infty} P\{(X_{n,r}^{(k)} - d)d^{-1/2} < x\} = (2\pi)^{-1/2} \int_{-x}^x \exp(-u^2/2) du, \quad -\infty < x < \infty,$$

where

$$d = d(n, r, k, p) = \frac{n^r}{r!} \binom{r}{k} [k(r-k)+1]^{r-k-2} p^a.$$

We omit a proof of this theorem because it follows similar lines as proofs of the respective theorems from [7] and [2].

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