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# LIMIT THEOREMS FOR ARRAYS OF MAXIMAL ORDER STATISTICS

#### BY

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Abstract. Let  $\{X, X_{nj}, 1 \leq j \leq m_n, n \geq 1\}$  be independent and identically distributed random variables with the Pareto distribution. Let  $X_{n(k)}$  be the k-th largest order statistic from the n-th row of our array. This paper establishes unusual limit theorems involving weighted sums for the sequence  $\{X_{n(k)}, n \geq 1\}$ .

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Consider independent and identically distributed random variables  $\{X, X_{nj}, 1 \le j \le m_n, n \ge 1\}$  with common density  $f_X(x) = px^{-p-1}I(x \ge 1)$ , where p > 0. Let  $X_{n(k)}$  be the k-th largest order statistic from each row of our array. Hence  $k \le m_n \to \infty$ . Therefore the density of  $X_{n(k)}$  is

$$f_{X_{n(k)}}(x) = \frac{p \cdot m_n!}{(m_n - k)! (k - 1)!} (1 - x^{-p})^{m_n - k} x^{-pk - 1} I(x \ge 1).$$

In this paper we will study limit theorems involving weighted sums of  $\{X_{n(k)}, n \ge 1\}$ . If pk > 1, then  $EX_{n(k)}$  is finite and the associated theorems are straightforward and unremarkable. If pk < 1, then these limit theorems fail to exist, see Theorem 4. The only interesting case is when pk = 1. Strange and unusual limit theorems occur when examining random variables that barely do or do not have a first moment. While this problem can be traced back to the St. Petersburg Game, some of the techniques in this paper can be traced back to Klass and Teicher [4] and Adler [1].

Our first theorem establishes a Weak Law, while our second result is a Generalized Law of the Iterated Logarithm and our third theorem is a closely related Strong Law. As usual, we define  $\lg x = \log(\max\{e, x\})$  and  $\lg_2 x = \lg(\lg x)$ . Also we use the constant C to denote a generic real number that is not necessarily the same in each appearance. THEOREM 1. If pk = 1 and  $\alpha > -1$ , then

$$\frac{\sum_{n=1}^{N} (n^{\alpha}/m_n^k) X_{n(k)}}{N^{\alpha+1} \lg N} \xrightarrow{P} \frac{1}{(\alpha+1) k!} \quad as \ N \to \infty.$$

Proof. Set  $a_n = n^{\alpha}/m_n^k$  and  $b_N = N^{\alpha+1} \lg N$ . From the Degenerate Convergence Theorem, which can be found on page 356 of Chow and Teicher [2], we have for all  $\varepsilon > 0$ 

$$\sum_{n=1}^{N} P\left\{X_{n(k)} > \varepsilon b_{N}/a_{n}\right\} < C \sum_{n=1}^{N} \frac{m_{n}!}{(m_{n}-k)!} \int_{\varepsilon b_{N}/a_{n}}^{\infty} (1-x^{-p})^{m_{n}-k} x^{-2} dx$$
$$< C \sum_{n=1}^{N} m_{n}^{k} \int_{\varepsilon b_{N}/a_{n}}^{\infty} x^{-2} dx < \frac{C}{b_{N}} \sum_{n=1}^{N} a_{n} m_{n}^{k}$$
$$= \frac{C \sum_{n=1}^{N} n^{\alpha}}{N^{\alpha+1} \lg N} < \frac{C N^{\alpha+1}}{N^{\alpha+1} \lg N} = \frac{C}{\lg N} \to 0.$$

Similarly, the variance term in the Degenerate Convergence Theorem is bounded above by

$$\sum_{n=1}^{N} \frac{a_n^2}{b_N^2} E X_{n(k)}^2 I(1 \le X_{n(k)} \le b_N/a_n) < \frac{C}{b_N^2} \sum_{n=1}^{N} a_n^2 m_n^k \int_{1}^{b_N/a_n} dx$$
$$< \frac{C \sum_{n=1}^{N} a_n m_n^k}{b_N} = \frac{C \sum_{n=1}^{N} n^\alpha}{N^{\alpha+1} \lg N} < \frac{C N^{\alpha+1}}{N^{\alpha+1} \lg N} = \frac{C}{\lg N} \to 0.$$

Next we observe where our truncated first moment is heading. From Gradshteyn and Ryzhik [3], page 4, we have

$$\sum_{j=1}^{m_n-k} \frac{\binom{m_n-k}{j}(-1)^j}{j} = -\sum_{j=1}^{m_n-k} \frac{1}{j}.$$

Thus

$$\sum_{n=1}^{N} E\left(\frac{a_n X_{n(k)}}{b_N} I\left(1 \leqslant X_{n(k)} \leqslant b_N/a_n\right)\right)$$
  
=  $\frac{1}{b_N k!} \sum_{n=1}^{N} \frac{a_n m_n!}{(m_n - k)!} \int_{1}^{b_N/a_n} (1 - x^{-p})^{m_n - k} x^{-1} dx$   
=  $\frac{1}{b_N k!} \sum_{n=1}^{N} \frac{a_n m_n!}{(m_n - k)!} \int_{1}^{b_N/a_n} \left[\frac{1}{x} + \sum_{j=1}^{m_n - k} \binom{m_n - k}{j}(-1)^j x^{-pj-1}\right] dx$   
=  $\frac{1}{b_N k!} \sum_{n=1}^{N} \frac{a_n m_n!}{(m_n - k)!} \left[ \lg b_N - \lg a_n + \sum_{j=1}^{m_n - k} \frac{\binom{m_n - k}{j}(-1)^j}{pj} \right]$ 

212

$$\begin{split} &+ \sum_{j=1}^{m_n-k} \frac{\binom{m_n-k}{j}(-1)^{j+1} a_n^{pj}}{pjb_N^{pj}} \\ &= \frac{1}{k! N^{\alpha+1} \lg N} \sum_{n=1}^N \frac{n^{\alpha} m_n!}{m_n^k (m_n-k)!} \left[ (\alpha+1) \lg N + \lg_2 N - \alpha \lg n + k \lg m_n \right] \\ &- k \sum_{j=1}^{m_n-k} \frac{1}{j} + \sum_{j=1}^{m_n-k} \frac{\binom{m_n-k}{j}(-1)^{j+1} (n^{\alpha}/m_n^k)^{pj}}{pj (N^{\alpha+1} \lg N)^{pj}} \\ &= \frac{(\alpha+1) \lg N + \lg_2 N}{k! N^{\alpha+1} \lg N} \sum_{n=1}^N \frac{n^{\alpha} m_n!}{m_n^k (m_n-k)!} - \frac{\alpha}{k! N^{\alpha+1} \lg N} \sum_{n=1}^N \frac{n^{\alpha} m_n! \lg n}{m_n^k (m_n-k)!} \\ &+ \frac{1}{(k-1)! N^{\alpha+1} \lg N} \sum_{n=1}^N \frac{n^{\alpha} m_n!}{m_n^k (m_n-k)!} \left[ \lg m_n - \sum_{j=1}^{m_n-k} \frac{1}{j} \right] \\ &+ \frac{1}{k! N^{\alpha+1} \lg N} \sum_{n=1}^N \frac{n^{\alpha} m_n!}{m_n^k (m_n-k)!} \sum_{j=1}^{m_n-k} \frac{\binom{m_n-k}{j}(-1)^{j+1} (n^{\alpha}/m_n^k)^{pj}}{pj (N^{\alpha+1} \lg N)^{pj}}. \end{split}$$

The first term is

$$\frac{(\alpha+1)\lg N + \lg_2 N}{k! N^{\alpha+1}\lg N} \sum_{n=1}^N \frac{n^{\alpha}m_n!}{m_n^k(m_n-k)!} \sim \frac{\alpha+1}{k! N^{\alpha+1}} \sum_{n=1}^N n^{\alpha} \to \frac{1}{k!}.$$

The second term is

$$\frac{-\alpha}{k! N^{\alpha+1} \lg N} \sum_{n=1}^{N} \frac{n^{\alpha} m_n! \lg n}{m_n^k (m_n - k)!} \sim \left(\frac{-\alpha}{k! N^{\alpha+1} \lg N}\right) \cdot \left(\sum_{n=1}^{N} n^{\alpha} \lg n\right)$$
$$\sim \left(\frac{-\alpha}{k! N^{\alpha+1} \lg N}\right) \cdot \left(\frac{N^{\alpha+1} \lg N}{\alpha+1}\right) = \frac{-\alpha}{(\alpha+1) k!}.$$

The third term is bounded above by

$$\left|\frac{1}{(k-1)! N^{\alpha+1} \lg N} \sum_{n=1}^{N} \frac{n^{\alpha} m_{n}!}{m_{n}^{k} (m_{n}-k)!} \left[ \lg m_{n} - \sum_{j=1}^{m_{n}-k} \frac{1}{j} \right] \right| < \frac{C}{N^{\alpha+1} \lg N} \sum_{n=1}^{N} n^{\alpha} < \frac{C}{\lg N} \to 0.$$

Finally, if we choose N large enough so that

$$\max_{1\leq n\leq N}\frac{n^{\alpha}}{N^{\alpha+1}\lg N}<\frac{1}{2^{1/p}},$$

then

$$\begin{aligned} \left| \frac{1}{k! N^{\alpha+1} \lg N} \sum_{n=1}^{N} \frac{n^{\alpha} m_{n}!}{m_{n}^{k} (m_{n}-k)!} \sum_{j=1}^{m_{n}-k} \frac{\binom{m_{n}-k}{j} (-1)^{j+1} (n^{\alpha}/m_{n}^{k})^{p_{j}}}{p_{j} (N^{\alpha+1} \lg N)^{p_{j}}} \right| \\ & < \frac{C}{N^{\alpha+1} \lg N} \sum_{n=1}^{N} n^{\alpha} \sum_{j=1}^{m_{n}} \frac{m_{n}^{j} (n^{\alpha})^{p_{j}}}{(m_{n}^{k} N^{\alpha+1} \lg N)^{p_{j}}} \\ & = \frac{C}{N^{\alpha+1} \lg N} \sum_{n=1}^{N} n^{\alpha} \sum_{j=1}^{m_{n}} \left( \frac{n^{\alpha}}{N^{\alpha+1} \lg N} \right)^{p_{j}} \\ & < \frac{C}{N^{\alpha+1} \lg N} \sum_{n=1}^{N} n^{\alpha} \sum_{j=1}^{\infty} \left( \frac{1}{2} \right)^{j} = \frac{C}{N^{\alpha+1} \lg N} \sum_{n=1}^{N} n^{\alpha} < \frac{C}{\lg N} \to 0. \end{aligned}$$

Therefore

$$\sum_{n=1}^{N} E\left(\frac{a_n X_{n(k)}}{b_N} I\left(1 \le X_{n(k)} \le b_N/a_n\right)\right) \to \frac{1}{k!} - \frac{\alpha}{(\alpha+1)\,k!} = \frac{1}{(\alpha+1)\,k!},$$

which completes the proof.

It is important to note that under the assumptions of Theorem 1 a Strong Law fails to hold. The ensuing result shows us the almost sure behaviour of the normalized partial sums observed in Theorem 1. This type of result is generally called a *Generalized Law of the Iterated Logarithm*.

THEOREM 2. If pk = 1 and  $\alpha > -1$ , then

$$\liminf_{N \to \infty} \frac{\sum_{n=1}^{N} (n^{\alpha}/m_n^k) X_{n(k)}}{N^{\alpha+1} \lg N} = \frac{1}{(\alpha+1) k!} \text{ almost surely}$$

and

$$\limsup_{N\to\infty}\frac{\sum_{n=1}^{N}(n^{\alpha}/m_{n}^{k})X_{n(k)}}{N^{\alpha+1}\lg N}=\infty \quad almost \ surely.$$

Proof. Let  $a_n = n^{\alpha}/m_n^k$ ,  $b_n = n^{\alpha+1} \lg n$  and  $c_n = b_n/a_n = nm_n^k \lg n$ . From Theorem 1 we can conclude that

$$\liminf_{N \to \infty} \frac{\sum_{n=1}^{N} (n^{\alpha}/m_n^k) X_{n(k)}}{N^{\alpha+1} \lg N} \leq \frac{1}{(\alpha+1)k!} \text{ almost surely.}$$

In order to establish the opposite inequality we note that

$$\frac{\sum_{n=1}^{N} a_n X_{n(k)}}{b_N} \ge \frac{\sum_{n=1}^{N} a_n X_{n(k)} I (1 \le X_{n(k)} \le nm_{n/2}^{k})}{b_N}$$

214

Limit theorems for order statistics

$$= \frac{\sum_{n=1}^{N} a_n [X_{n(k)} I(1 \le X_{n(k)} \le nm_n^k) - EX_{n(k)} I(1 \le X_{n(k)} \le nm_n^k)]}{b_N} + \frac{\sum_{n=1}^{N} a_n EX_{n(k)} I(1 \le X_{n(k)} \le nm_n^k)}{b_N}.$$

The first term vanishes almost surely by the usual Khintchine-Kolmogorov Convergence Theorem (see Chow and Teicher [2]) and Kronecker's lemma since

$$\sum_{n=1}^{\infty} c_n^{-2} E X_{n(k)}^2 I(1 \le X_{n(k)} \le nm_n^k) < C \sum_{n=1}^{\infty} \frac{m_n^k}{[nm_n^k \lg n]^2} \int_1^{nm_n^k} dx - C \sum_{n=1}^{\infty} \frac{nm_n^{2k}}{n^2 m_n^{2k} (\lg n)^2} = C \sum_{n=1}^{\infty} \frac{1}{n (\lg n)^2} < \infty.$$

As for the second term

$$\begin{split} EX_{n(k)} I\left(1 \leqslant X_{n(k)} \leqslant nm_{n}^{k}\right) &\sim \frac{m_{n}^{k}}{k!} \int_{1}^{nm_{n}^{k}} (1 - x^{-p})^{m_{n}-k} x^{-1} dx \\ &= \frac{m_{n}^{k}}{k!} \int_{1}^{nm_{n}^{k}} \left[ \frac{1}{x} + \sum_{j=1}^{m_{n}-k} \binom{m_{n}-k}{j} (-1)^{j} x^{-pj-1} \right] dx \\ &= \frac{m_{n}^{k}}{k!} \left[ \lg\left(nm_{n}^{k}\right) + \sum_{j=1}^{m_{n}-k} \frac{\binom{m_{n}-k}{j} (-1)^{j}}{pj} + \sum_{j=1}^{m_{n}-k} \frac{\binom{m_{n}-k}{j} (-1)^{j+1}}{pj (nm_{n}^{k})^{pj}} \right] \\ &= \frac{m_{n}^{k}}{k!} \left[ \lg n + k \lg m_{n} - k \sum_{j=1}^{m_{n}-k} \frac{1}{j} + \sum_{j=1}^{m_{n}-k} \frac{\binom{m_{n}-k}{j} (-1)^{j+1}}{pj (nm_{n}^{k})^{pj}} \right] \\ &= \frac{m_{n}^{k}}{k!} \left[ \lg n + k \left[ \lg m_{n} - \sum_{j=1}^{m_{n}-k} \frac{1}{j} \right] + k \sum_{j=1}^{m_{n}-k} \frac{\binom{m_{n}-k}{j} (-1)^{j+1}}{j (nm_{n}^{k})^{pj}} \right] \\ &\sim \frac{m_{n}^{k} \lg n}{k!} \end{split}$$

since

$$\lg m_n - \sum_{j=1}^{m_n - k} \frac{1}{j} = O(1)$$

and

$$\sum_{j=1}^{m_n-k} \frac{\binom{m_n-k}{j}(-1)^{j+1}}{j(nm_n^k)^{p_j}} \bigg| < \sum_{j=1}^{m_n} \frac{m_n^j}{(nm_n^k)^{p_j}} = \sum_{j=1}^{m_n} \left(\frac{1}{n^p}\right)^j = O(1).$$

A. Adler

Thus

$$\frac{\sum_{n=1}^{N} a_n E X_{n(k)} I\left(1 \leqslant X_{n(k)} \leqslant nm_n^k\right)}{b_N} \sim \frac{\sum_{n=1}^{N} (n^\alpha/m_n^k) \cdot \left((m_n^k \lg n)/k!\right)}{N^{\alpha+1} \lg N}$$
$$= \frac{\sum_{n=1}^{N} n^\alpha \lg n}{k! N^{\alpha+1} \lg N} \to \frac{1}{(\alpha+1)k!},$$

whence

$$\liminf_{N \to \infty} \frac{\sum_{n=1}^{N} (n^{\alpha}/m_n^k) X_{n(k)}}{N^{\alpha+1} \lg N} \ge \frac{1}{(\alpha+1) k!} \text{ almost surely,}$$

which leads us to equality for the lower limit. As for the upper limit, we let M > 0. Then

$$\sum_{n=1}^{\infty} P\left\{X_{n(k)} > Mc_{n}\right\} = \sum_{n=1}^{\infty} \frac{m_{n}!}{(m_{n}-k)!k!} \int_{Mc_{n}}^{\infty} (1-x^{-p})^{m_{n}-k} x^{-2} dx$$

$$= \frac{1}{k!} \sum_{n=1}^{\infty} \frac{m_{n}!}{(m_{n}-k)!} \sum_{j=0}^{m_{n}-k} \binom{m_{n}-k}{j} (-1)^{j} \int_{Mc_{n}}^{\infty} x^{-pj-2} dx$$

$$= \frac{1}{k!} \sum_{n=1}^{\infty} \frac{m_{n}!}{(m_{n}-k)!} \sum_{j=0}^{m_{n}-k} \binom{m_{n}-k}{j} \frac{(-1)^{j}}{(pj+1)(Mc_{n})^{pj+1}}$$

$$= \frac{1}{M(k-1)!} \sum_{n=1}^{\infty} \frac{m_{n}!}{c_{n}(m_{n}-k)!} \left[ \frac{1}{k} + \sum_{j=1}^{m_{n}-k} \binom{m_{n}-k}{j} \frac{(-1)^{j}}{(j+k)(Mc_{n})^{pj}} \right]$$

$$> C \sum_{n=1}^{\infty} \frac{m_{n}!}{c_{n}(m_{n}-k)!}$$

since

$$\left|\sum_{j=1}^{m_n-k} \binom{m_n-k}{j} \frac{(-1)^j}{(j+k)(Mc_n)^{pj}}\right| < \sum_{j=1}^{m_n} \frac{m_n^j}{(Mc_n)^{pj}} = \sum_{j=1}^{m_n} \frac{m_n^j}{M^{pj} [nm_n^k \lg n]^{pj}} < \sum_{j=1}^{\infty} \left(\frac{1}{[Mn \lg n]^p}\right)^j = \frac{1}{[Mn \lg n]^p - 1} \to 0.$$

Hence

$$\sum_{n=1}^{\infty} P\{X_{n(k)} > Mc_n\} > C \sum_{n=1}^{\infty} \frac{m_n!}{c_n(m_n-k)!}$$
$$= C \sum_{n=1}^{\infty} \frac{m_n!}{nm_n^k \lg n(m_n-k)!} > C \sum_{n=1}^{\infty} \frac{1}{n \lg n} = \infty,$$

216

which implies that

$$\limsup_{n \to \infty} \frac{a_n X_{n(k)}}{b_n} = \infty \text{ almost surely,}$$

which in turn allows us to conclude that

$$\limsup_{N\to\infty}\frac{\sum_{n=1}^N a_n X_{n(k)}}{b_N}=\infty \text{ almost surely,}$$

completing this proof.

What proves to be quite interesting is that if we let  $\alpha = -1$  in Theorems 1 and 2, then not only does a Weak Law, but also a Strong Law exists. This next result is our *Strong Law of Large Numbers*. Notice that the norming sequence is different, we need an extra logarithm term.

THEOREM 3. If pk = 1 and  $\alpha > -2$ , then

$$\lim_{N \to \infty} \frac{\sum_{n=1}^{N} ((\lg n)^{\alpha} / nm_n^k) X_{n(k)}}{(\lg N)^{\alpha+2}} = \frac{1}{(\alpha+2) k!} \text{ almost surely.}$$

Proof. Let  $a_n = (\lg n)^{\alpha}/(nm_n^k)$ ,  $b_n = (\lg n)^{\alpha+2}$  and  $c_n = b_n/a_n = nm_n^k(\lg n)^2$ . We partition our sum into the following three terms:

$$\frac{\sum_{n=1}^{N} a_n X_{n(k)}}{b_N} = \frac{\sum_{n=1}^{N} a_n [X_{n(k)} I(1 \le X_{n(k)} \le c_n) - EX_{n(k)} I(1 \le X_{n(k)} \le c_n)]}{b_N} + \frac{\sum_{n=1}^{N} a_n X_{n(k)} I(X_{n(k)} > c_n)}{b_N} + \frac{\sum_{n=1}^{N} a_n EX_{n(k)} I(1 \le X_{n(k)} \le c_n)}{b_N}$$

The first term vanishes almost surely since

$$\sum_{n=1}^{\infty} c_n^{-2} E X_{n(k)}^2 I(1 \le X_{n(k)} \le c_n) < C \sum_{n=1}^{\infty} \frac{m_n^k}{c_n^2} \int_{1}^{c_n} dx$$
$$< C \sum_{n=1}^{\infty} \frac{m_n^k}{c_n} = C \sum_{n=1}^{\infty} \frac{1}{n(\lg n)^2} < \infty.$$

The second term vanishes almost surely since

$$\sum_{n=1}^{\infty} P\left\{X_{n(k)} > c_n\right\} < C \sum_{n=1}^{\infty} m_n^k \int_{c_n}^{\infty} x^{-2} \, dx = C \sum_{n=1}^{\infty} \frac{m_n^k}{c_n} = C \sum_{n=1}^{\infty} \frac{1}{n (\lg n)^2} < \infty.$$

2 - PAMS 22.2

As for the third term

$$\begin{split} EX_{n(k)} I\left(1 \leqslant X_{n(k)} \leqslant c_{n}\right) &= \frac{p \cdot m_{n}!}{(m_{n} - k)! (k - 1)!} \int_{1}^{c_{n}} (1 - x^{-p})^{m_{n} - k} x^{-1} dx \\ &= \frac{p \cdot m_{n}!}{(m_{n} - k)! (k - 1)!} \int_{1}^{c_{n}} \left[\frac{1}{x} + \sum_{j=1}^{m_{n} - k} \binom{m_{n} - k}{j} (-1)^{j} x^{-pj - 1}\right] dx \\ &= \frac{p \cdot m_{n}!}{(m_{n} - k)! (k - 1)!} \left[\lg c_{n} + \sum_{j=1}^{m_{n} - k} \binom{m_{n} - k}{j} (-1)^{j} + \sum_{j=1}^{m_{n} - k} \binom{m_{n} - k}{j} (-1)^{j+1} - \frac{pjc_{n}^{pj}}{jc_{n}^{pj}}\right] \\ &= \frac{m_{n}!}{(m_{n} - k)! k!} \left[k \lg m_{n} + \lg n + 2 \lg_{2} n - k \sum_{j=1}^{m_{n} - k} \frac{1}{j} + \sum_{j=1}^{m_{n} - k} \binom{m_{n} - k}{j} (-1)^{j+1} - \frac{pjc_{n}^{pj}}{jc_{n}^{pj}}\right] \\ &= \frac{m_{n}!}{(m_{n} - k)! k!} \left[\lg n + 2 \lg_{2} n + k \left[\lg m_{n} - \sum_{j=1}^{m_{n} - k} \frac{1}{j}\right] + \sum_{j=1}^{m_{n} - k} \binom{m_{n} - k}{j} (-1)^{j+1} - \frac{m_{n} - k}{j} - \frac{m_{n} - k}{j} (-1)^{j+1} - \frac{m_{n} - k}{j} (-1)^{j} - \frac{m_{n} - k}{j} - \frac{m_{n$$

since  $\lg m_n - \sum_{j=1}^{m_n} j^{-1} = O(1)$ , and if we let  $n(\lg n)^2 > 2^{1/p}$ , then

$$\left| \sum_{j=1}^{m_n-k} \frac{\binom{m_n-k}{j} (-1)^{j+1}}{p j c_n^{pj}} \right| < \sum_{j=1}^{m_n-k} \frac{m_n^j}{c_n^{pj}} < \sum_{j=1}^{\infty} \frac{m_n^j}{[nm_n^k (\lg n)^2]^{pj}} \\ = \sum_{j=1}^{\infty} \left[ \frac{1}{[n(\lg n)^2]^p} \right]^j < \sum_{j=1}^{\infty} \left( \frac{1}{2} \right)^j = 1$$

Thus we are able to show that the final term converges to the desired result, i.e.,

$$\frac{\sum_{n=1}^{N} a_n E X_{n(k)} I(1 \leq X_{n(k)} \leq c_n)}{b_N} \sim \frac{\sum_{n=1}^{N} ((\lg n)^{\alpha} / nm_n^k) \cdot ((m_n^k \lg n) / k!)}{(\lg N)^{\alpha+2}}$$
$$= \frac{\sum_{n=1}^{N} (\lg n)^{\alpha+1} / n}{k! (\lg N)^{\alpha+2}} \to \frac{1}{(\alpha+2) k!},$$

which completes the proof.

If pk < 1, then under rather mild conditions no interesting results can occur.

THEOREM 4. If pk < 1 and  $a_n$  and  $b_N$  are positive constants, where (1)  $\max_{1 \le n \le N} m_n a_n^p = o(b_N^p),$  then the only finite limit of our normalized sums is zero, i.e.,

$$\frac{\sum_{n=1}^{N} a_n X_{n(k)}}{b_N} \xrightarrow{P} 0 \quad as \ N \to \infty.$$

**Proof.** Let  $w_n$  be the median from the density of  $X_{n(k)}$ . Thus

$$\frac{1}{2} = \frac{p \cdot m_n!}{(m_n - k)! (k - 1)!} \int_{w_n}^{\infty} (1 - x^{-p})^{m_n - k} x^{-pk - 1} dx$$
$$< \frac{p \cdot m_n^k}{(k - 1)!} \int_{w_n}^{\infty} x^{-pk - 1} dx = \frac{m_n^k}{k! w_n^{pk}}.$$

Hence, we can conclude that  $w_n < Cm_n^{1/p}$  and from (1) we have

$$\frac{\max_{1\leq n\leq N}a_nw_n}{b_N} < \frac{C\max_{1\leq n\leq N}a_nm_n^{1/p}}{b_N} \to 0.$$

Assuming that a Weak Law holds we have, by the Degenerate Convergence Theorem,

$$0 \leftarrow \sum_{n=1}^{N} P\left\{X_{n(k)} > b_{N}/a_{n}\right\} = \frac{p}{(k-1)!} \sum_{n=1}^{N} \frac{m_{n}!}{(m_{n}-k)!} \int_{b_{N}/a_{n}}^{\infty} (1-x^{-p})^{m_{n}-k} x^{-pk-1} dx$$
$$= \frac{p}{(k-1)!} \sum_{n=1}^{N} \frac{m_{n}!}{(m_{n}-k)!} \sum_{j=0}^{m_{n}-k} \binom{m_{n}-k}{j} (-1)^{j} \int_{b_{N}/a_{n}}^{\infty} x^{-p(j+k)-1} dx$$
$$= \frac{p}{(k-1)!} \sum_{n=1}^{N} \frac{m_{n}!}{(m_{n}-k)!} \sum_{j=0}^{m_{n}-k} \frac{\binom{m_{n}-k}{j} (-1)^{j} (a_{n}/b_{N})^{p(j+k)}}{p(j+k)}$$
$$= \frac{1}{(k-1)!} \sum_{n=1}^{N} \frac{m_{n}!}{(m_{n}-k)!} \left(\frac{a_{n}}{b_{N}}\right)^{pk} \left[\frac{1}{k} + \sum_{j=1}^{m_{n}-k} \frac{\binom{m_{n}-k}{j} (-1)^{j}}{j+k} \left(\frac{a_{n}}{b_{N}}\right)^{pj}\right]$$
$$> C \sum_{n=1}^{N} \frac{m_{n}!}{(m_{n}-k)!} \left(\frac{a_{n}}{b_{N}}\right)^{pk}$$

since, for if we select N large enough so that  $m_n a_n^p < \varepsilon b_N^p$  for all  $1 \le n \le N$  and  $0 < \varepsilon < 1/2$ , it follows that

$$\begin{aligned} &\sum_{j=1}^{m_n-k} \frac{\binom{m_n-k}{j}(-1)^j}{j+k} \binom{a_n}{b_N}^{pj} \middle| &< \sum_{j=1}^{m_n-k} \binom{m_n-k}{j} \binom{a_n}{b_N}^{pj} \\ &< \sum_{j=1}^{m_n-k} (m_n-k)^j \binom{a_n}{b_N}^{pj} < \sum_{j=1}^{\infty} m_n^j \binom{a_n}{b_N}^{pj} = \sum_{j=1}^{\infty} \binom{m_n a_n^p}{b_N^p}^j < \sum_{j=1}^{\infty} \varepsilon^j < 2\varepsilon. \end{aligned}$$

Therefore

$$\sum_{n=1}^{N} \frac{m_n!}{(m_n-k)!} \left(\frac{a_n}{b_N}\right)^{pk} \to 0.$$

Then, by once again utilizing the Degenerate Convergence Theorem, the limit of our normalized partial sum is zero since

$$\sum_{n=1}^{N} \frac{a_n}{b_N} E X_{n(k)} I(1 \leq X_{n(k)} \leq b_N/a_n) < \frac{C}{b_N} \sum_{n=1}^{N} \frac{a_n m_n!}{(m_n - k)!} \int_{1}^{b_N/a_n} x^{-pk} dx$$
$$< \frac{C}{b_N} \sum_{n=1}^{N} \frac{a_n m_n!}{(m_n - k)!} \left(\frac{b_N}{a_n}\right)^{-pk+1} = C \sum_{n=1}^{N} \frac{m_n!}{(m_n - k)!} \left(\frac{a_n}{b_N}\right)^{pk} \to 0,$$

which completes the proof.

In conclusion there are two final comments. The first is that (1) is quite mild. Note that (1) holds for all the selected constants,  $a_n$  and  $b_N$ , in our first three theorems. The other comment is that if pk > 1, then a Strong Laws exists since  $EX_{n(k)}$  is finite.

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