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CONDITIONED LIMIT THEOREMS FOR FUNCTIONS OF THE AVERAGE OF LI.D. RANDOM VARIABLES

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Abstract. Let $\{\xi_k, k \ge 1\}$ be a sequence of i.i.d. random variables with $E\xi_1 = 0$, $0 < E\xi_1^2 = \sigma^2 < \infty$. Form the random walk $\{S_n, n \ge 0\}$ by setting $S_0 = 0$, $S_n = \xi_1 + \ldots + \xi_n$, $n \ge 1$. Let T denote the hitting time of the set $(-\infty, 0]$ by the random walk. Put $X_n(t) = S_{[nt]}/\sigma \sqrt{n}$, $0 \le t \le 1$. Let h be a real-valued, right-continuous function on R, having left limits, with h(0) = 1, and continuous at 0. For $\beta > 0$ we define the map H_n : $D[0, 1] \rightarrow D[0, 1]$ by $H_n(f) = fh(n^{-\beta} f)$, $f \in D[0, 1]$, $n \ge 1$. Put $Y_n = H_n(X_n)$. This note deals with the asymptotic behaviour of Y_n conditioned on [T > n]. Moreover, we investigate the asymptotic behaviour in the question when n is replaced by N_n , where $\{N_n, n \ge 1\}$ is a sequence of positive integer-valued random variables.

1. Introduction. Let $\{\xi_k, k \ge 1\}$ be a sequence of independent, identically distributed random variables with $E\xi_1 = 0$, $0 < E\xi_1^2 = \sigma^2 < \infty$, and let $\{N_m, m \ge 0\}$, $N_0 = 0$ a.s., be a sequence of positive integer-valued random variables. Form the random walk $\{S_n, n \ge 0\}$ by setting $S_0 = 0$ and $S_n = \xi_1 + \ldots + \xi_n$, $n \ge 1$. Define the random function X_n by

$$X_n(t) = S_{[nt]} / \sigma_{n} / n, \quad 0 \le t \le 1,$$

where [x] is the greatest integer in x. Next let T be the hitting time of the set $(-\infty, 0]$ by the random walk,

$$T = \inf \{n > 0; S_n \leq 0\},\$$

where the infimum of the empty set is taken to be $+\infty$. Let *h* be a real-valued, right-continuous function on **R**, having left limits, with h(0) = 1 and continuous at 0. Let $D \equiv D[0, 1]$ be the space of real-valued, right-conti-

nuous functions on [0, 1] having left limits. For $\beta > 0$ we define the map H_n : $D[0, 1] \rightarrow D[0, 1]$ by

$$H_n(f) = fh(n^{-\beta}f), \quad f \in D[0, 1], \ n \ge 1.$$

Put $Y_n = H_n(X_n)$. The aim of this note is to give a functional central limit theorem for the random function Y_n , conditioned on [T > n].

To be more specific we assume that $\{\xi_k, k \ge 1\}$ are the coordinate functions defined on the product space (Ω, \mathscr{A}, P) . If $\Lambda_n = \{T > n\}$, then we let $(\Lambda_n, \Lambda_n \cap \mathscr{A}, P_n)$ be the trace of (Ω, \mathscr{A}, P) on Λ_n , where $\Lambda_n \cap \mathscr{A}$ $= \{\Lambda_n \cap F, F \in \mathscr{A}\}$ and $P_n[A] = P[A]/P[\Lambda_n]$ for $A \in \Lambda_n \cap \mathscr{A}$. Let \mathscr{D} be the σ -field of Borel sets on D, generated by the open sets of the Skorohod \mathscr{J}_1 topology. Let $D_+ = \{x \in D; x \ge 0\}$, and $\mathscr{D}_+ = D_+ \cap \mathscr{D}$. The measurable mappings $X_n^+, Y_n^+: (\Lambda_n, \Lambda_n \cap \mathscr{A}) \to (D_+, \mathscr{D}_+)$ are defined by

$$X_n^+(\cdot, \omega) = S_{[n\cdot]}(\omega) / \sigma \sqrt{n}, \quad \omega \in \Lambda_n,$$

and

$$Y_n^+(\cdot, \omega) = H_n(X_n^+(\cdot, \omega)), \quad \omega \in \Lambda_n.$$

The random function X_n^+ induces a probability measure (p. m.) μ_n^+ on \mathcal{D}_+ : for $A \in \mathcal{D}_+$ we have

$$\mu_n^+(A) = \mathcal{P}_n[X_n^+ \in A] = \mathcal{P}[X_n^+ \in A]/\mathcal{P}[A_n] = \mathcal{P}[X_n \in A | A_n].$$

Iglehart [5], Theorem (3.4) (see also for this result Bolthausen [2]), has proved that $X_n^+ \Rightarrow W^+$, $n \to \infty$, i.e. $\mu_n^+ \Rightarrow \mu^+$, the p.m. of a Brownian meander W^+ , the symbol \Rightarrow means weak convergence. Alternatively, we write $(X_n|\Lambda_n) \Rightarrow W^+$, $n \to \infty$, for this result. The random function Y_n^+ induces a p.m. on \mathcal{D}_+ : for $A \in D_+$ we have

$$\overline{\mu}_n^+(A) = \mathbb{P}_n[Y_n^+ \in A] = \mathbb{P}[Y_n^+ \in A]/\mathbb{P}[A_n].$$

The main result of this note is that $\bar{\mu}_n^+ \Rightarrow \mu^+$, $n \to \infty$. Moreover, we investigate the asymptotic behaviour in the question when *n* is replaced by N_n , where $\{N_n, n \ge 0\}$, $N_0 = 0$ a.s., is a sequence of positive integer-valued random variables.

We shall apply the following result of Doney [3]: LEMMA 1. For v > 0

(1) $\lim_{n \to \infty} \Pr\left[\max_{0 \le k \le T} S_k / \sigma \sqrt{n} \le v | \Lambda_n\right] = 1 - v^{-1} \sqrt{\frac{\pi}{2}} + 2 \sum_{k=1}^{\infty} \exp\left(-2k^2 v^2\right).$

We need in the sequel the following lemmas: LEMMA 2. Let $\beta > 0$. For $\delta > 0$ we have

$$\limsup_{n\to\infty} \mathbb{P}\left[\sup_{0\leqslant t\leqslant 1}|X_n(t)|>\delta|\Lambda_n\right]\leqslant \sqrt{\frac{\pi}{2}}\delta^{-1},$$

(2)

and (3)

$$\lim_{n \to \infty} \mathbb{P}\left[\sup_{0 \le t \le 1} |n^{-\beta} X_n(t)| > \delta |\Lambda_n\right] = 0.$$

Proof. We have

$$\mathbb{P}\left[\sup_{0 \le t \le 1} |X_n(t)| > \delta |A_n\right] = \mathbb{P}\left[\sup_{0 \le t \le 1} S_{[nt]} / \sqrt{n\sigma} > \delta |A_n\right]$$

 $\leq \mathbb{P}\left[\max_{0\leq k\leq T}S_k>\sigma\sqrt{n}\,\delta|\Lambda_n\right].$

Hence, (2) follows from Lemma 1.

We now prove (3). For any given $\varepsilon > 0$ and $\delta > 0$ we have

 $\limsup_{n\to\infty} \mathbb{P}\left[\sup_{0\leq t\leq 1}|X_n(t)|>\delta n^{\beta}|A_n\right]$

 $\leq \limsup_{n \to \infty} \mathbb{P}\left[\sup_{0 \leq t \leq 1} |X_n(t)| > \frac{\delta}{\varepsilon} \middle| A_n\right] \leq \sqrt{\frac{\pi}{2}} \frac{\varepsilon}{\delta},$

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which implies

$$\lim_{n \to \infty} \mathbb{P}\left[\sup_{0 \le t \le 1} |n^{-\beta} X_n(t)| > \delta |\Lambda_n\right] = 0$$

and completes the proof of Lemma 2.

LEMMA 3. Let $\beta > 0$. For any $\delta > 0$

(4)
$$\lim_{n\to\infty} \mathbb{P}\left[\sup_{0\leq t\leq 1}|Y_n(t)-X_n(t)|>\delta|\Lambda_n\right]=0.$$

Proof. Let $\varepsilon > 0$. Since h is continuous at the point x = 0, it follows that there exists an $\eta > 0$ such that

(5)
$$|x| \leq \eta \Rightarrow |h(x) - 1| < \varepsilon.$$

Then, for any given $\delta > 0$, we have

$$\begin{split} & \mathbb{P}\left[\sup_{0 \leq t \leq 1} |Y_n(t) - X_n(t)| > \delta |\Lambda_n\right] \\ & \leq \mathbb{P}\left[\sup_{0 \leq t \leq 1} |X_n(t)| \left| 1 - h\left(n^{-\beta} X_n(t)\right) \right| > \delta, \sup_{0 \leq t \leq 1} |n^{-\beta} X_n(t)| \leq \eta |\Lambda_n\right] \\ & + \mathbb{P}\left[\sup_{0 \leq t \leq 1} |n^{-\beta} X_n(t)| > \eta |\Lambda_n\right] \end{split}$$

$$\mathbb{P}\left[\sup_{0 \le t \le 1} |X_n(t)| > \frac{\delta}{\varepsilon} \middle| \Lambda_n \right] + \mathbb{P}\left[\sup_{0 \le t \le 1} |n^{-\beta} X_n(t)| > \eta |\Lambda_n\right].$$

Hence, by Lemma 2, we get

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$$\limsup_{n\to\infty} \mathbb{P}\left[\sup_{0\leq t\leq 1}|Y_n(t)-X_n(t)|>\delta|\Lambda_n\right] \leq \sqrt{\frac{\pi}{2}}\frac{\varepsilon}{\delta}$$

which proves (4).

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Let d denote the metric on D ([1], p. 111). We can rewrite Lemma 3 in the following form:

(6)
$$\lim_{n\to\infty} \mathbb{P}\left[d(X_n, Y_n) > \delta | \Lambda_n\right] = 0.$$

THEOREM 1. Under the assumptions of this note we have

(7)
$$Y_n^+ = (Y_n | A_n) \Rightarrow W^+, \quad n \to \infty.$$

Proof. Assertion (7) follows from (5) and a result of Iglehart $(X_n^+ \Rightarrow W^+, n \to \infty)$.

Denote by \mathscr{G} the class of all continuous functions g differentiable at the point x = 0 with $g'(0) \neq 0$.

COROLLARY 1. For every $g \in \mathcal{G}$

(8)
$$((\sqrt{n}/g'(0)\sigma)(g(S_{[n-1]}/n)-g(0))|\Lambda_n) \Rightarrow W^+, \quad n \to \infty,$$

holds true.

Proof. Putting

$$h(x) = \begin{cases} \frac{g(x) - g(0)}{xg'(0)} & \text{if } x \neq 0, \\ 1 & \text{if } x = 0, \end{cases}$$

we can write the left-hand side of (8) in the following form:

$$(\sqrt{n/g'(0)}\,\sigma)(g(S_{[n\cdot]}/n)-g(0))=X_nh(n^{-1/2}X_n).$$

It is easy to verify that h satisfies our assumptions. Putting in Theorem 1 that $\beta = 1/2$, we obtain (8).

2. Random partial sum processes. In this section we are interested in the asymptotic behaviour of

$$Y_{N_m}^+ = (Y_{N_m} | \Lambda_{N_m}), \quad m \ge 1,$$

where $\{N_m, m \ge 0\}$, $N_0 = 0$ a.s., is a sequence of positive integer-valued random variables.

We now need the following extension of Lemma 1:

LEMMA 4. Let $\{\xi_k, k \ge 1\}$ be a sequence of i.i.d. random variables with $E\xi_1 = 0, \ 0 < E\xi_1^2 = \sigma^2 < \infty$. Suppose that $\{N_m, m \ge 1\}$ is a sequence of positive integer-valued random variables, and $\{\alpha_m, m \ge 1\}$ is a sequence of positive numbers with $\alpha_m \to \infty, \ m \to \infty$, such that, for any given $\varepsilon > 0$,

9)
$$\mathbb{P}\left[|N_m/\alpha_m - 1| \ge \varepsilon\right] = o\left(1/\sqrt{\alpha_m}\right).$$

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Then, for v > 0,

(10)
$$\mathbb{P}\left[\max_{0 \le k \le T} \frac{S_k}{\sigma \sqrt{N_m}} \le v | T > N_m\right] = 1 - v^{-1} \sqrt{\frac{\pi}{2}} + 2 \sum_{k=1}^{\infty} \exp(-2k^2 v^2).$$

Proof. From Theorem 3.7, [6], we have

$$P[T > n] = P[S_1 > 0, ..., S_n > 0] \sim c/\sqrt{n}, \quad n \to \infty,$$

where

$$c = \exp \left\{ \sum_{k=1}^{\infty} (1/k) (1/2 - P[S_k > 0]) \right\}.$$

Hence, by (9), for $1 > \varepsilon > 0$ we have

(11)
$$\frac{c}{\sqrt{[(1-\varepsilon)\alpha_m]}} - o\left(\frac{1}{\sqrt{\alpha_m}}\right) \leq \Pr[T > N_m]$$
$$\leq \frac{c}{\sqrt{[(1+\varepsilon)\alpha_m]}} + o\left(\frac{1}{\sqrt{\alpha_m}}\right), \quad m \to \infty$$

Put $A_m = \{k; (1-\varepsilon)\alpha_m \le k \le (1+\varepsilon)\alpha_m\}$, and let A_m^c denote the complement of A_m . Then, by (9) and (11), we have

$$\mathbf{P}[N_m \in A_m^c] / \mathbf{P}[T > N_m] \to 0, \quad m \to \infty.$$

Hence, we can write

$$P\left[\max_{0 \le k \le T} S_k / \sigma \sqrt{N_m} \le v | T > N_m\right]$$

=
$$P\left[\max_{0 \le k \le T} S_k / \sigma \sqrt{N_m} \le v, T > N_m, N_m \in A_m\right] / P\left[T > N_m\right] + P\left[\max_{0 \le k \le T} S_k / \sigma \sqrt{N_m} \le v, T > N_m, N_m \in A_m^c\right] / P\left[T > N_m\right].$$

We get thus the following estimate:

$$\begin{split} & \mathbb{P}\left[\max_{0 \leq k \leq T} S_{k} / \sigma \sqrt{\left[(1-\varepsilon)\alpha_{m}\right]} \leq v, \ T > \\ & \left[(1+\varepsilon)\alpha_{m}\right]\right] / \mathbb{P}\left[T > N_{m}\right] - \mathbb{P}\left[N_{m} \in A_{m}^{c}\right] / \mathbb{P}\left[T > N_{m}\right] \\ & \leq \mathbb{P}\left[\max_{0 \leq k \leq T} S_{k} / \sigma \sqrt{N_{m}} \leq v | \ T > N_{m}\right] \leq \mathbb{P}\left[\max_{0 \leq k \leq T} S_{k} / \sigma \sqrt{\left[(1+\varepsilon)\alpha_{m}\right]} \\ & \leq v, \ T > \left[(1-\varepsilon)\alpha_{m}\right]\right] / \mathbb{P}\left[T > N_{m}\right] + \mathbb{P}\left[N_{m} \in A_{m}^{c}\right] / \mathbb{P}\left[T > N_{m}\right]. \end{split}$$

Therefore, by (1) and letting $m \to \infty$ and next $\varepsilon \to 0$, we get (10). From Lemmas 4 and 2 one can get the following

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LEMMA 5. Let $\{\xi_k, k \ge 1\}$ and $\{N_m, m \ge 0\}$ be as in Lemma 4 and let $\beta > 0$. For $\delta > 0$ we have

(12)
$$\limsup_{m \to \infty} \mathbb{P}\left[\sup_{0 \le t \le 1} |X_{N_m}(t)| > \delta | T > N_m\right] \le \sqrt{\frac{\pi}{2}} \delta^{-1}$$

and

(13)
$$\lim_{m \to \infty} \Pr[\sup_{0 \le t \le 1} |N_m^{-\beta} X_{N_m}(t)| > \delta |T > N_m] = 0.$$

LEMMA 6. Under the assumptions of this note, for any $\delta > 0$,

(14)
$$\lim_{m \to \infty} \mathbb{P}\left[\sup_{0 \le t \le 1} |Y_{N_m}(t) - X_{N_m}(t)| > \delta |T > N_m\right] = 0,$$

where $Y_{N_m}(t) = X_{N_m}(t) h(N_m^{-\beta} X_{N_m}(t)).$

Note now that Theorem 3 of [7] implies $(X_{N_m} | T > N_m) \Rightarrow W^+, m \to \infty$. Thus, by (14), we get the following theorem:

THEOREM 2. Let $\{\xi_k, k \ge 1\}$ be a sequence of i.i.d. random variables with $E\xi_1 = 0, 0 < E\xi_1^2 = \sigma^2 < \infty$. Suppose that $\{N_m, m \ge 1\}$ is a sequence of positive integer-valued random variables, and $\{\alpha_m, m \ge 1\}$ is a sequence of positive numbers with $\alpha_m \to \infty, m \to \infty$, such that, for any given $\varepsilon > 0$,

(15)
$$\mathbf{P}[|N_m/\alpha_m - 1| \ge \varepsilon] = o(1/\sqrt{\alpha_m}).$$

Then .

(16)

$$(Y_{N_m}|T>N_m) \Rightarrow W^+, \quad m \to \infty.$$

COROLLARY 2. For every $g \in \mathcal{G}$,

(17)
$$((\sqrt{N_m/g'(0)\sigma})(g(S_{[N_m]}/N_m)-g(0))|T>N_m) \Rightarrow W^+, \quad m \to \infty.$$

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