

Class 9

- Data, model and inference for multiple regression

Data for Multiple Regression

- Y_i is the response variable
- $X_{i1}, X_{i2}, \dots, X_{ip-1}$ are $p-1$ explanatory variables for cases $i = 1$ to n

Multiple Regression Model

- $Y_i = \beta_0 + \beta_1 X_{i1} + \beta_2 X_{i2} + \dots + \beta_{p-1} X_{ip-1} + \xi_i$
- Y_i is the value of the response variable for the i^{th} case
- β_0 is the intercept
- $\beta_1, \beta_2, \dots, \beta_{p-1}$ are the regression coefficients for the explanatory variables

Multiple Regression Model (2)

- X_{ik} is the value of the k^{th} explanatory variable for the i^{th} case
- ξ_i are independent normally distributed random errors with mean 0 and variance σ^2

Many interesting special cases

- $Y_i = \beta_0 + \beta_1 X_i + \beta_2 X_i^2 + \dots + \beta_{p-1} X_i^{p-1} + \xi_i$
- X s can be *indicator* or *dummy* variables with 0 and 1 (or any other two distinct numbers) as possible values
- Interactions
- $Y_i = \beta_0 + \beta_1 X_{i1} + \beta_2 X_{i2} + \beta_3 X_{i1} X_{i2} + \xi_i$

Multiple Regression Parameters

- β_0 the intercept
- $\beta_1, \beta_2, \dots, \beta_{p-1}$ the regression coefficients for the explanatory variables
- σ^2 the variance of the error term

Model in Matrix Form

$$\begin{matrix} \mathbf{Y} & = & \mathbf{X} & \boldsymbol{\beta} + & \boldsymbol{\xi} \\ \text{nx1} & & \text{nxp} & \text{px1} & \text{nx1} \end{matrix}$$

$$\boldsymbol{\xi} \sim N(0, \sigma^2 \mathbf{I})$$

$$\mathbf{Y} \sim N(\mathbf{X}\boldsymbol{\beta}, \sigma^2 \mathbf{I})$$

Least Squares

$$\mathbf{Y} = \mathbf{X}\boldsymbol{\beta} + \boldsymbol{\xi}$$

$$\min(\mathbf{Y} - \mathbf{X}\mathbf{b})'(\mathbf{Y} - \mathbf{X}\mathbf{b})$$

$$\mathbf{X}'\mathbf{X}\mathbf{b} = \mathbf{X}'\mathbf{Y}$$

Least Squares Solution

$$\mathbf{b} = (\mathbf{X}'\mathbf{X})^{-1} \mathbf{X}'\mathbf{Y}$$

Fitted (predicted) values

$$\begin{aligned} \hat{\mathbf{Y}} &= \mathbf{X}\mathbf{b} = \mathbf{X}(\mathbf{X}'\mathbf{X})^{-1} \mathbf{X}'\mathbf{Y} \\ &= \mathbf{H}\mathbf{Y} \end{aligned}$$

Residuals

$$\mathbf{e} = \mathbf{Y} - \hat{\mathbf{Y}}$$

$$= \mathbf{Y} - \mathbf{H}\mathbf{Y}$$

$$= (\mathbf{I} - \mathbf{H})\mathbf{Y}$$

$\mathbf{I} - \mathbf{H}$ is symmetric and idempotent i.e.

$$(\mathbf{I} - \mathbf{H})(\mathbf{I} - \mathbf{H}) = (\mathbf{I} - \mathbf{H})$$

Covariance Matrix of residuals

- $\text{Cov}(\mathbf{e}) = \sigma^2(\mathbf{I} - \mathbf{H})(\mathbf{I} - \mathbf{H})' = \sigma^2(\mathbf{I} - \mathbf{H})$
- So,
- $\text{Var}(\mathbf{e}_i) = \sigma^2(1 - h_{ii})$
- $h_{ij} = \mathbf{X}'_i(\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}_j$
- $\mathbf{X}'_i = (1, \mathbf{X}_{i1}, \dots, \mathbf{X}_{i(p-1)})$
- Residuals are usually correlated
- $\text{Cov}(\mathbf{e}_i, \mathbf{e}_j) = -\sigma^2 h_{ij}$

Estimation of σ

$$\begin{aligned} s^2 &= \frac{\mathbf{e}'\mathbf{e}}{n - p} \\ &= \frac{(\mathbf{Y} - \mathbf{X}\mathbf{b})'(\mathbf{Y} - \mathbf{X}\mathbf{b})}{n - p} \end{aligned}$$

$$= \frac{SSE}{dfe} = MSE$$

$$s = \sqrt{s^2} = \text{Root MSE}$$

Distribution of b

- $b = (X'X)^{-1}X'Y$
- $Y \sim N(X\beta, \sigma^2 I)$
- $E(b) = ((X'X)^{-1}X')X\beta = \beta$
- $Cov(b) = \sigma^2 ((X'X)^{-1}X') ((X'X)^{-1}X')'$
 $= \sigma^2 (X'X)^{-1}$

Estimation of variance of b

- $b \sim N(\beta, \sigma^2 (X'X)^{-1})$
- $\sigma^2 (X'X)^{-1}$
- Is estimated by
- $s^2 (X'X)^{-1}$

ANOVA Table

- To organize arithmetic
- Sources of variation are
 - Model
 - Error or Residual
 - Total
- SS and df add
 - $SSM + SSE = SST$
 - $dfM + dfE = dfT$

SS

$$SSM = \sum_{i=1}^n (\hat{Y}_i - \bar{Y})^2$$

$$SSE = \sum_{i=1}^n (Y_i - \hat{Y}_i)^2$$

$$SST = \sum_{i=1}^n (Y_i - \bar{Y})^2$$

df

$$df \text{ M} = p - 1$$

$$df \text{ E} = n - p$$

$$df \text{ T} = n - 1$$

Mean Squares

$$MSM = SSM/dfM$$

$$MSE = SSE/dfE$$

$$MST = SST/dfT$$

Mean Squares (2)

$$\text{MSM} = \sum_{i=1}^n (\hat{Y}_i - \bar{Y})^2 / (p-1)$$

$$\text{MSE} = \sum_{i=1}^n (Y_i - \hat{Y}_i)^2 / (n-p)$$

$$\text{MST} = \sum_{i=1}^n (Y_i - \bar{Y})^2 / (n-1)$$

ANOVA Table

Source	SS	df	MS	F
Model	SSM	dfM	MSM	MSM/MSE
Error	SSE	dfE	MSE	
Total	SST	dfT (MST)		

ANOVA F test

- $H_0: \beta_1 = \beta_2 = \dots \beta_{p-1} = 0$
- $H_a: \beta_k \neq 0$, for at least one $k=1, \dots, p-1$
- Under H_0 , $F \sim F(p-1, n-p)$
- Reject H_0 if F is large, use P value

Study of CS students

- Study of computer science majors at Purdue
- Large drop out rate
- Can we find predictors of success ?
- Predictors must be available at time of entry into program

Data available

- GPA after three semesters
- High school math grades
- High school science grades
- High school English grades
- SAT Math
- SAT Verbal
- Gender (of interest for other reasons)

Example

```

Data a1;
  infile '.../csdata.dat';
  input id gpa
        hsm hss hse
        satm satv genderm1;
proc reg data=a1;
  model gpa=hsm hss hse;
run;

```

CS ANOVA Table

Source	DF	Sum of Squares	Mean Square	F
Model	3	27.71	9.23	18.86
Error	220	107.75	0.48	
Total	223	135.46		

Hypothesis Tested by F

- $H_0: \beta_1 = \beta_2 = \dots \beta_{p-1} = 0$
- $F = \text{MSM}/\text{MSE}$
- Reject H_0 if the P value is $\leq .05$

What do we conclude?

Sum of Source	Mean Square	F	P
Model	9.23744	18.86	<.0001
Error	0.48977		

R^2

- The squared multiple regression correlation (R^2) gives the proportion of variation in the response variable explained by the explanatory variables included in the model
- It is usually expressed as a percent
- It is sometimes called the coefficient of multiple determination

R^2 (2)

- $R^2 = \text{SSM}/\text{SST}$, the proportion of variation explained
- $R^2 = 1 - (\text{SSE}/\text{SST})$, 1 – the proportion of variation not explained
- $F = [(R^2)/(p-1)] / [(1 - R^2)/(n-p)]$
- For the CS data $R^2 = 27.71/135.46 = 0.205$

- The P-value for the F significance test tells us one of the following:
 - there is no evidence to conclude that *any* of our explanatory variables can help us to model the response variable using this kind of model ($P \geq .05$)
 - one or more of the explanatory variables in our model *is* potentially useful for predicting the response variable in a linear model ($P \leq .05$)

Inference for individual regression coefficients

- $\mathbf{b} \sim N(\boldsymbol{\beta}, \sigma^2 (\mathbf{X}'\mathbf{X})^{-1})$
- $\mathbf{S}^2_{\mathbf{b}} = \mathbf{s}^2 (\mathbf{X}'\mathbf{X})^{-1}$
- $\mathbf{s}^2(\mathbf{b}_i) = \mathbf{S}^2_{\mathbf{b}}(i,i)$
- CI: $\mathbf{b}_i \pm t_c \mathbf{s}(\mathbf{b}_i)$, where $t_c = t(.975, n-p)$
- Significance test for $H_{0i}: \beta_i = 0$ uses the test statistic $t = \mathbf{b}_i / \mathbf{s}(\mathbf{b}_i)$, $df = dfE = n-p$, and the P-value computed from the $t(n-p)$ distribution

CS Data results

Variable	DF	Parameter	Standard	t Value	Pr > t
		Estimate	Error		
Intercept	1	0.58988	0.29424	2.00	0.0462
hsm	1	0.16857	0.03549	4.75	<.0001
hss	1	0.03432	0.03756	0.91	0.3619
hse	1	0.04510	0.03870	1.17	0.2451

Example

- Dwaine Studios operates portrait studios in 21 cities
- Y is sales
- X_1 is number of persons aged 16 and under
- X_2 is per capita disposable income
- $n = 21$ cities

Check the data

```
data a1;
infile '.../ch06fi05.txt';
input young income sales;
proc print data=a1;
run;
```

Print

Obs	young	income	sales
1	68.5	16.7	174.4
2	45.2	16.8	164.4
3	91.3	18.2	244.2
4	47.8	16.3	154.6
5	46.9	17.3	181.6

Proc Reg

```
proc reg data=a1;
model sales=young income;
run;
```

Output

Source	DF	Sum of Sq	Mean Sq	F	Pr > F
Model	2	24015	12008	99.	<.0001
Error	18	2180	121		
C Tot	20	26196			

Output (2)

Root MSE 11.00 R-Sq 0.916
 Dep Mean 181.9 Adj R-Sq 0.907
 Coeff Var 6.0

Output (3)

Var	DF	Par Est	St Err	t	P
Int	1	-68.8	60.0	-1.15	0.2663
young	1	1.454	0.21	6.87	<.0001
income	1	9.365	4.06	2.30	0.0333

CLB option

```
proc reg data=a1;
  model sales=young income/clb;
run;
```

Output (4)

95% Confidence Limits

Int	-194.94	57.23
young	1.00	1.89
income	0.82	17.90

Estimation of $E(Y_h)$

- X_h is now a vector
- $(1, X_{h1}, X_{h2}, \dots, X_{h1})'$
- We want a point estimate and a confidence interval for the subpopulation mean corresponding to X_h

Theory for $E(Y_h)$

$$E(Y_h) = \mu_h = X'_h \beta$$

$$\hat{\mu}_h = X'_h b$$

$$\sigma^2(\hat{\mu}_h) = X'_h \sum_b X_h = \sigma^2 X'_h (X'X)^{-1} X_h$$

$$s^2(\hat{\mu}_h) = s^2 X'_h (X'X)^{-1} X_h$$

$$CI : \hat{\mu}_h \pm s(\hat{\mu}_h) t_{(0.975, n-p)}$$

Estimation of $E(Y_h)$ (CLM)

```
proc reg data=a1;
  model sales=young income/clm;
  id young income;
run;
```

$E(Y_h)$ CLM Output

			DV	Pre	SE
Obs	young	inco	sal	Val	Pre
1	68.5	16.7	174	187	3.8
2	45.2	16.8	164	154	3.5
3	91.3	18.2	244	234	4.5
4	47.8	16.3	154	153	3.2

$E(Y_h)$ CLM Output (2)

Obs	95% CL	Mean
1	179.1146	195.2536
2	146.7591	161.6998
3	224.7569	244.0358
4	146.5361	160.1210

Prediction of Y_h

- X_h is now a vector
- $(1, X_{h1}, X_{h2}, \dots, X_{h1})'$
- We want a prediction for Y_h with an interval that expresses the uncertainty in our prediction

Theory for Y_h

$$Y_h = X'_h \beta + \xi$$

$$\hat{Y}_h = \hat{\mu}_h = X'_h b$$

$$\begin{aligned} \sigma^2(pred) &= Var(\hat{Y}_h - Y_h) = Var \hat{Y}_h + \sigma^2 \\ &= \sigma^2 (1 + X'_h (X'X)^{-1} X_h) \end{aligned}$$

$$s^2(pred) = s^2 (1 + X'_h (X'X)^{-1} X_h)$$

$$CI : \hat{\mu}_h \pm s(pred) t_{(0.975, n-p)}$$

Prediction of Y_h (CLI)

```
proc reg data=a1;  
  model sales=young income/cli;  
  id young income;  
run;
```

Prediction Intervals Output

Obs	95% CL Predict	
1	162.6910	211.6772
2	129.9271	178.5317
3	209.3421	259.4506
4	129.2260	177.4311