

Lecture 6

- Data, model and inference for multiple regression

Data for Multiple Regression

- Y_i is the response variable
- $X_{i1}, X_{i2}, \dots, X_{ip-1}$ are $p-1$ explanatory variables for cases $i = 1$ to n

Multiple Regression Model

- $Y_i = \beta_0 + \beta_1 X_{i1} + \beta_2 X_{i2} + \dots + \beta_{p-1} X_{ip-1} + \xi_i$
- Y_i is the value of the response variable for the i^{th} case
- β_0 is the intercept
- $\beta_1, \beta_2, \dots, \beta_{p-1}$ are the regression coefficients for the explanatory variables

Multiple Regression Model (2)

- X_{ik} is the value of the k^{th} explanatory variable for the i^{th} case
- ξ_i are independent normally distributed random errors with mean 0 and variance σ^2

Many interesting special cases

- $Y_i = \beta_0 + \beta_1 X_i + \beta_2 X_i^2 + \dots + \beta_{p-1} X_i^{p-1} + \xi_i$
- Xs can be *indicator* or *dummy* variables with 0 and 1 (or any other two distinct numbers) as possible values
- Interactions
- $Y_i = \beta_0 + \beta_1 X_{i1} + \beta_2 X_{i2} + \beta_3 X_{i1} X_{i2} + \xi_i$

Multiple Regression Parameters

- β_0 the intercept
- $\beta_1, \beta_2, \dots, \beta_{p-1}$ the regression coefficients for the explanatory variables
- σ^2 the variance of the error term

Model in Matrix Form

$$\begin{array}{cccccc} \mathbf{Y} & = & \mathbf{X} & \boldsymbol{\beta} & + & \boldsymbol{\xi} \\ & nx1 & nxp & px1 & & nx1 \end{array}$$

$$\boldsymbol{\xi} \sim N(0, \sigma^2 \mathbf{I})$$

$$\mathbf{Y} \sim N(\mathbf{X}\boldsymbol{\beta}, \sigma^2 \mathbf{I})$$

Least Squares

$$\mathbf{Y} = \mathbf{X}\boldsymbol{\beta} + \boldsymbol{\xi}$$

$$\min(\mathbf{Y} - \mathbf{X}\mathbf{b})'(\mathbf{Y} - \mathbf{X}\mathbf{b})$$

$$\mathbf{X}'\mathbf{X}\mathbf{b} = \mathbf{X}'\mathbf{Y}$$

Least Squares Solution

$$\mathbf{b} = (\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'\mathbf{Y}$$

Fitted (predicted) values

$$\begin{aligned} \hat{\mathbf{Y}} &= \mathbf{X}\mathbf{b} = \mathbf{X}(\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'\mathbf{Y} \\ &= \mathbf{HY} \end{aligned}$$

Residuals

$$\mathbf{e} = \mathbf{Y} - \hat{\mathbf{Y}}$$

$$= \mathbf{Y} - \mathbf{HY}$$

$$= (\mathbf{I} - \mathbf{H})\mathbf{Y}$$

$\mathbf{I} - \mathbf{H}$ is symmetric and idempotent i.e.
 $(\mathbf{I} - \mathbf{H})(\mathbf{I} - \mathbf{H}) = (\mathbf{I} - \mathbf{H})$

Covariance Matrix of residuals

- $\text{Cov}(\mathbf{e}) = \sigma^2(\mathbf{I} - \mathbf{H})(\mathbf{I} - \mathbf{H})' = \sigma^2(\mathbf{I} - \mathbf{H})$
- So,
- $\text{Var}(e_i) = \sigma^2(1 - h_{ii})$
- $h_{ii} = \mathbf{x}_i'(\mathbf{X}'\mathbf{X})^{-1}\mathbf{x}_i$
- $\mathbf{x}_i' = (1, x_{i1}, \dots, x_{i(p-1)})$
- Residuals are usually correlated
- $\text{Cov}(e_i, e_j) = -\sigma^2 h_{ij}$

Estimation of σ

$$\begin{aligned} s^2 &= \frac{\mathbf{e}'\mathbf{e}}{n-p} \\ &= \frac{(\mathbf{Y} - \mathbf{X}\mathbf{b})'(\mathbf{Y} - \mathbf{X}\mathbf{b})}{n-p} \\ &= \frac{SSE}{df\mathbf{e}} = MSE \end{aligned}$$

$$s = \sqrt{s^2} = \text{Root MSE}$$

Distribution of \mathbf{b}

- $\mathbf{b} = (\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'\mathbf{Y}$
- $\mathbf{Y} \sim \mathcal{N}(\mathbf{X}\beta, \sigma^2\mathbf{I})$
- $E(\mathbf{b}) = ((\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}')\mathbf{X}\beta = \beta$
- $Cov(\mathbf{b}) = \sigma^2 ((\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}') ((\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}')'$
 $= \sigma^2 (\mathbf{X}'\mathbf{X})^{-1}$

Estimation of variance of \mathbf{b}

- $\mathbf{b} \sim \mathcal{N}(\beta, \sigma^2 (\mathbf{X}'\mathbf{X})^{-1})$
- $\sigma^2 (\mathbf{X}'\mathbf{X})^{-1}$
- Is estimated by
- $s^2 (\mathbf{X}'\mathbf{X})^{-1}$

ANOVA Table

- To organize arithmetic
- Sources of variation are
 - Model
 - Error
 - Total
- SS and df add
 - $SSM + SSE = SST$
 - $dfM + dfE = dfT$

SS

$$SSM = \sum_{i=1}^n (\hat{Y}_i - \bar{Y})^2$$

$$SSE = \sum_{i=1}^n (Y_i - \hat{Y}_i)^2$$

$$SST = \sum_{i=1}^n (Y_i - \bar{Y})^2$$

df

$$df M = p - 1$$

$$df E = n - p$$

$$df T = n - 1$$

Mean Squares

$$MSM = SSM/dfM$$

$$MSE = SSE/dfE$$

$$MST = SST/dfT$$

Mean Squares (2)

$$MSM = \sum_{i=1}^n (\hat{Y}_i - \bar{Y})^2 / (p-1)$$

$$MSE = \sum_{i=1}^n (Y_i - \hat{Y}_i)^2 / (n-p)$$

$$MST = \sum_{i=1}^n (Y_i - \bar{Y})^2 / (n-1)$$

ANOVA Table

Source	SS	df	MS	F
Model	SSM	dfM	MSM	MSM/MSE
Error	SSE	dfE	MSE	
Total	SST	dfT	(MST)	

ANOVA F test

- $H_0: \beta_1 = \beta_2 = \dots = \beta_{p-1} = 0$
- $H_a: \beta_k \neq 0$, for at least one $k=1, \dots, p-1$
- Under H_0 , $F \sim F(p-1, n-p)$
- Reject H_0 if F is large, use P value

Study of CS students

- Study of computer science majors at Purdue
- Large drop out rate
- Can we find predictors of success
- Predictors must be available at time of entry into program

Data available

- GPA after three semesters
- High school math grades
- High school science grades
- High school English grades
- SAT Math
- SAT Verbal
- Gender (of interest for other reasons)

Example

```
cs<-read.table('csdata.dat',
col.names=c("id", "gpa", "hsm",
"hss", "hse", "satm", "satv",
"gen"));
reg1<-lm(gpa~hsm+hss+hse, cs);
Anova(reg1);
summary(reg1);
```

CS ANOVA Table

	Df	Sum	Mean	F	Pr(>F)
hsm	1	25.81	25.8	52.7	6.6e-12
hss	1	1.24	1.23	2.5	0.1134
hse	1	0.67	0.67	1.4	0.2451
Res	220	107.7	0.49		

F-stat: 18.86 on 3 and 220 DF
p-value: 6.359e-11

Hypothesis Tested by F

- $H_0: \beta_1 = \beta_2 = \dots = \beta_{p-1} = 0$
- $F = MSM/MSE$
- Reject H_0 if the P value is $\leq .05$
- What do we conclude ?

R²

- The squared multiple regression correlation (R^2) gives the proportion of variation in the response variable explained by the explanatory variables included in the model
- It is usually expressed as a percent
- It is sometimes called the coefficient of multiple determination

R² (2)

- $R^2 = SSM/SST$, the proportion of variation explained
- $R^2 = 1 - (SSE/SST)$, 1 – the proportion of variation not explained
- $F = [(R^2)/(p-1)] / [(1-R^2)/(n-p)]$

- The P-value for the F significance test tells us one of the following:
 - there is no evidence to conclude that *any* of our explanatory variables can help us to model the response variable using this kind of model ($P \geq .05$)
 - one or more of the explanatory variables in our model *is* potentially useful for predicting the response variable in a linear model ($P \leq .05$)

Stat 512 Class 14

- Review multiple linear regression
 - data
 - Model
- Inference for multiple regression (continued)
- Diagnostics and remedies

Data for Multiple Regression

- Y_i is the response variable
- $X_{i1}, X_{i2}, \dots, X_{ip-1}$ are $p-1$ explanatory variables for cases $i = 1$ to n
- $Y_i, X_{i1}, X_{i2}, \dots, X_{ip-1}$ is the data for case i , where $i = 1$ to n
- Y/X is the data

Multiple Regression Model

- $Y_i = \beta_0 + \beta_1 X_{i1} + \beta_2 X_{i2} + \dots + \beta_{p-1} X_{ip-1} + \xi_i$
- Y_i is the value of the response variable for the i^{th} case
- β_0 is the intercept
- $\beta_1, \beta_2, \dots, \beta_{p-1}$ are the regression coefficients for the explanatory variables

Multiple Regression Model (2)

- X_{ik} is the value of the k^{th} explanatory variable for the i^{th} case
- ξ_i are independent normally distributed random errors with mean 0 and variance σ^2

Model in Matrix Form

$$\begin{matrix} Y \\ nx1 \end{matrix} = \begin{matrix} X \\ nxp \end{matrix} \begin{matrix} \beta \\ px1 \end{matrix} + \begin{matrix} \xi \\ nx1 \end{matrix}$$

$$\xi \sim N(0, \sigma^2 I)$$

$$Y \sim N(X\beta, \sigma^2 I)$$

Least Squares Solution

$$\mathbf{b} = (\mathbf{X}' \mathbf{X})^{-1} \mathbf{X}' \mathbf{Y}$$

Estimation of σ^2

$$\begin{aligned} s^2 &= \text{MSE} \\ s &= \text{Root MSE} \end{aligned}$$

ANOVA F test

- $H_0: \beta_1 = \beta_2 = \dots = \beta_{p-1} = 0$
- $H_a: \beta_k \neq 0$, for at least one $k=1, \dots, p-1$
- Under H_0 , $F \sim F(p-1, n-p)$
- Reject H_0 if F is large, using P value we reject if $P \leq 0.05$

R²

- $R^2 = SSM/SST$, the proportion of variation explained by the explanatory variables

Inference for individual regression coefficients

- $\hat{\beta} \sim N(\beta, \sigma^2(X'X)^{-1})$
- $S^2_{\hat{\beta}} = s^2(X'X)^{-1}$
- $s^2(\hat{\beta}_i) = S^2_{\hat{\beta}}(i,i)$
- CI: $\hat{\beta}_i \pm t^* s(\hat{\beta}_i)$, where $t^* = t(.975, n-p)$
- Significance test for $H_{0i}: \beta_{i1} = 0$ uses the test statistic $t = \hat{\beta}_i / s(\hat{\beta}_i)$, $df = dfE = n-p$, and the P-value computed from the $t(n-p)$ distribution

Example

- Dwaine Studios operates portrait studios in 21 cities
- Y is sales
- X_1 is number of persons aged 16 and under
- X_2 is per capita disposable income
- $n = 21$ cities

R code

```
dwst<-read.table('ch06fi05.txt',
col.names=c("young", "income",
"sales"));
reg<-lm(sales~young+income,
dwst);
summary(reg)
```

	Est	Std	t	p-val
Int	-68.86	60.02	-1.15	0.2663
young	1.45	0.21	6.87	2e-06
income	9.37	4.06	2.31	0.0333

Residual standard error: 11.01
on 18 degrees of freedom
Multiple R-squared: 0.9167,
Adjusted R-squared: 0.9075
F-statistic: 99.1 on 2 and 18
DF, p-value: 1.921e-10

```

confint(reg)

      2.5 %    97.5 %
Int   -194.9480130 57.233867
young   1.0096226 1.899497
income   0.8274411 17.903560

```

Estimation of $E(Y_h)$

- X_h is now a vector
- $(1, X_{h1}, X_{h2}, \dots, X_{h1})'$
- We want a point estimate and a confidence interval for the subpopulation mean corresponding to X_h

Theory for $E(Y_h)$

$$\begin{aligned}
E(Y_h) &= \mu_h = X'_h \beta \\
\hat{\mu}_h &= X'_h b \\
\sigma^2(\hat{\mu}_h) &= X'_h \sum_b X_h = \sigma^2 X'_h (X'X)^{-1} X_h \\
s^2(\hat{\mu}_h) &= s^2 X'_h (X'X)^{-1} X_h \\
CI : \hat{\mu}_h &\pm s(\hat{\mu}_h) t_{(0.975, n-p)}
\end{aligned}$$

Estimation of $E(Y_h)$ (CLM)

```

predict.lm(reg,
interval='confidence');

```

$E(Y_h)$ CI Output

	fit	lwr	upr
1	187.1841	179.1146	195.2536
2	154.2294	146.7591	161.6998
3	234.3963	224.7569	244.0358
4	153.3285	146.5361	160.1210
5	161.3849	152.0778	170.6921

Prediction of Y_h

- X_h is now a vector
- $(1, X_{h1}, X_{h2}, \dots, X_{h1})'$
- We want a prediction for Y_h with an interval that expresses the uncertainty in our prediction

Theory for Y_h

$$Y_h = X'_h \beta + \xi$$

$$\hat{Y}_h = \hat{\mu}_h = X'_h b$$

$$\begin{aligned}\sigma^2(pred) &= \text{Var}(\hat{Y}_h - Y_h) = \text{Var} \hat{Y}_h + \sigma^2 \\ &= \sigma^2(1 + X'_h (X'X)^{-1} X_h)\end{aligned}$$

$$s^2(pred) = s^2(1 + X'_h (X'X)^{-1} X_h)$$

$$CI : \hat{\mu}_h \pm s(pred) t_{(0.975, n-p)}$$

Prediction of Y_h (PI)

```
predict.lm(reg,  
interval='prediction');
```

Prediction Intervals Output

	fit	lwr	upr
1	187.1841	162.6910	211.6772
2	154.2294	129.9271	178.5317
3	234.3963	209.3421	259.4506
4	153.3285	129.2260	177.4311
5	161.3849	136.4566	186.3132

Diagnostics

- Look at the distribution of each variable
- Look at the relationship between pairs of variables
- Plot the residuals versus
 - Each explanatory variable
 - Time

Diagnostics (2)

- Are the residuals approximately normal?
 - Look at a histogram
 - Normal quantile plot
- Is the variance constant?
 - Plot the squared residuals vs anything that might be related to the variance (e.g. residuals vs predicted)

Remedial measures

- Transformations such as Box-Cox
- Analyze without outliers

Scatter Plot Matrix

```
pairs(~gpa+satm+satv,cs)
```

